



Prevalence of human papillomavirus detection in ovarian cancer: a meta-analysis

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Abstract

We conducted a meta-analysis of published data to update and estimate the prevalence of HPV in ovarian cancer. A comprehensive literature search was performed according to the PRISMA guidelines. Eligible articles published from 1989 until 2020 by searching Web of Sciences, Pubmed, Embase, and the Cochrane Library Central databases were gathered. A pooled estimation of HPV prevalence with a 95% confidence interval (CI) was calculated based on a random effect model. Quantitative assessment of heterogeneity was explored using Cochrane test and I^2 . Additionally, publication bias, sensitivity, meta-regression, and subgroup analyses were also performed. Twenty-nine studies involving 2280 patients with ovarian cancer were included. The statistical heterogeneity was high ($I^2 = 88%$, $P < 0.0001$). The pooled prevalence of HPV in ovarian cancer cases was 15.9% (95% CI, 11–22). In subgroup analyses, the highest prevalence of HPV was reported by studies from Asia (30.9%; 95% CI, 20–44) and Eastern Europe (29.3%; 95% CI, 4.4–78). Furthermore, the most frequently detected HPV genotype was HPV16 (54%; 95% CI, 27.9–55), followed by HPV18 (23.2%; 95% CI, 18.8–28.2). Our meta-analysis suggests a great difference in the prevalence of HPV detected in ovarian cancer by different studies, which is not seen in strongly HPV-associated cancers such as cervical cancer. However, the prevalence varied markedly by geographic region. Considering the substantial heterogeneity found, more studies with control groups and precise assays measuring HPV mRNA expression are needed to further evaluate the link and causative aetiology between HPV and ovarian cancer.

Keywords HPV prevalence · Ovarian malignancy · Worldwide · Subgroup analysis · Etiologic agent

Background

Ovarian cancer remains the most life-threatening malignancy of the female genital tract mainly because of the lack of early clinical symptoms and early detection (FIGO stage I–II) [1]. The stage of ovarian cancer is an important prognostic factor at diagnosis [2, 3]. This malignancy is typically diagnosed at a

late stage (FIGO stage III–IV) with a 5 year-survival rate reaching 20% whereas it is 80–90% for patients with localized tumor (stage I) [4, 5]. Despite the considerable advances in highlighting risk factors, the pathogenesis and aetiology of ovarian cancer are still unclear [1]. However, an increased risk of developing ovarian cancer has been strongly correlated to genetic mutations (BRCA1 or BRCA2 genes), family history of ovarian, breast, or colon cancer, age, postmenopausal hormonal therapy use, infertility, and nulliparity [4]. Besides these findings, a current theory, supported by epidemiological data, hypothesizes that persistent viral infection and chronic inflammation may play a role in the carcinogenesis of ovarian cancer [2]. Thus, it is hypothesized that viral infection may contribute to ovarian cancer.

HPV has been identified as an etiological agent of numerous proliferative epithelial lesions in the skin and various mucosal sites, including the lower genital tract (of both sexes) and diverse sites of the oropharyngeal area and the upper aerodigestive tract [6]. However, its role in the development

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of cancers in the upper genital tract, such as endometrial and ovarian cancer is still under debate [7]. In fact, studies carried out so far failed to associate firmly the presence of HPV with the occurrence of these malignancies. Some reported a positive correlation while others indicated negative results [1, 6, 8]. These discordant findings might be due to the fact that most studies on this issue were observational reports without including control groups [2, 9]. Furthermore, controversial results might be due to variation in the geographic distribution of HPV, and the technique used to identify HPV infection.

Two systematic reviews and meta-analyses have studied the association between HPV and ovarian cancer by investigating the prevalence of HPV in ovarian cancer tissues [2, 10]. The results, however, were conflicting. Besides, data from the Middle East had not been included. As more studies concerning HPV detection in ovarian cancer have been published recently, we undertook the present meta-analysis to update and better define this relationship.

Material and methods

Study identification and data extraction

Two investigators (S.C. and A.A.) performed a systematic literature research independently by using Pubmed, Embase, and Cochrane Library Central between 1989 until 2020, using the following Medical Subject Heading terms (MeSH) and text words: “Human papillomavirus”, “HPV”, “ovarian malignancies”, “ovarian neoplasm”, and “ovarian cancer”. The investigators extracted independently data from identified studies; additional studies were retrieved and reviewed. In the case of discrepancy, the decision on inclusion/exclusion was made by discussion. The meta-analysis was performed in agreement with PRISMA criteria.

Data extraction was carried out, in an Excel sheet, to record details of the first author, country of publication, method of detection, histological type, type of specimen, sample size, HPV genotype, and numbers of HPV-positive and negative patients.

Study selection

The following criteria had to be met to include studies in this meta-analysis: (1) observational studies published between 1989 until 2020 with data on the association of HPV in ovarian cancer and including at least 5 cases, (2) written in English, and (3) published as a full peer-reviewed article. The exclusion criteria were the following: (1) studies where a serological test is used to detect HPV, (2) studies not meeting the inclusion criteria, and (3) studies limited to animals.

Statistical analysis

The pooled prevalence of HPV in ovarian cancer and 95% confidence interval (CI) were calculated by using the Mantel-Haenszel method of DerSimonian and Laird method (random-effects or fixed-effects model). The heterogeneity was measured by the Cochran Q test ($P < 0.10$ demonstrates a high level of heterogeneity) [11, 12]. Moreover, the rate of inconsistency (I^2) was also calculated (values of I^2 from 50 to 75% correspond to moderate to high degrees of heterogeneity, respectively) [13]. If the heterogeneity was not substantial, the pooled rate with a 95% confidence interval (95%CI) was calculated using the fixed-effect model. When statistically significant heterogeneity was found ($I^2 > 50$), the pooled rate with 95% CI was calculated based on the random-effects model. Then, a subgroup analysis was employed to evaluate the influence of several factors on the overall results. A sensitivity analysis was also performed, by excluding one study each time, to assess the impact of each study on the strength and stability of our results. Meta-regression analysis was used to examine the association of the geographical distribution of the studies, specimen type and detection methods with the prevalence of HPV. The statistical analyses were performed by using the Comprehensive Meta-analysis software, version 3 (Englewood, USA).

Publication bias

The publication bias was evaluated visually by constructing a Funnel plot. It was created by plotting the log prevalence of HPV vs the standard error. Their symmetry was evaluated by Egger’s regression test and the Begg and Mazumdar adjusted rank correlation test, all P -values were set on two sides, those less than 0.05 were deemed statistically significant.

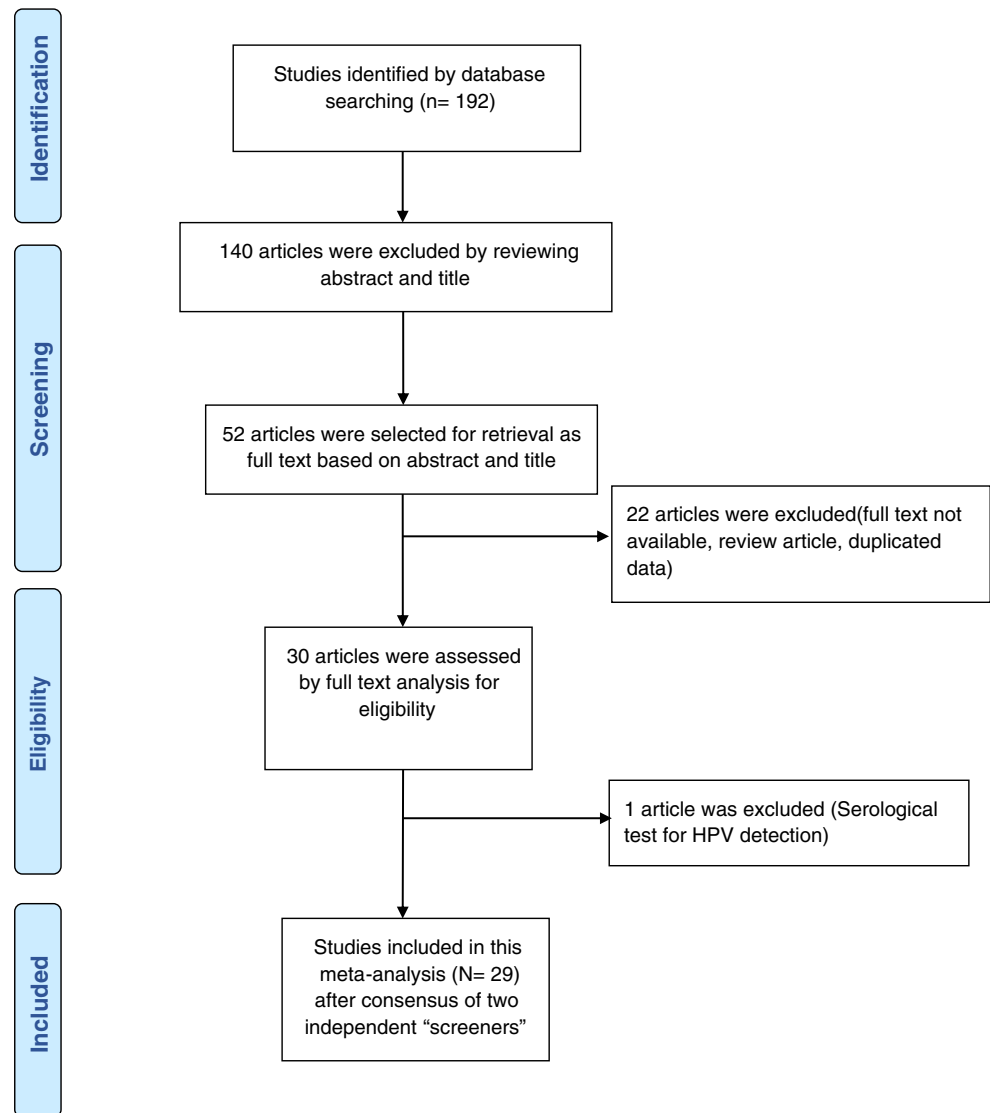
Results

Description of studies

Of the 192 articles initially identified by the investigators, 29 studies were included in this meta-analysis. The flowchart explaining the study selection is displayed in Fig. 1. The main characteristics of the studies are presented in Supplementary information 1. In 19/29 studies, the histological types were mentioned.

A total of 2280 patients were included in this meta-analysis. There were 6 studies from North America [14–19], 8 studies from the Middle East [1, 8, 20–25], 7 studies were conducted in Asia [7, 26–31] and the remaining 8 studies were reported from Europe [6, 9, 32–37]. As for the type of specimen utilized, Formalin-Fixed Paraffin-Embedded (FPPE) tissues were used in 19 studies [1, 6, 8, 9, 15, 17, 18, 20–27, 29,

Fig. 1 Flow diagram of the studies selected for this meta-analysis



30, 33, 36], in the remaining studies, frozen [7, 14, 19, 31, 32, 34, 35, 37] and fresh tissues [16, 28] were utilized.

The most used technique for the detection of HPV was PCR (25 studies) using the L1 consensus primers either by a single-step PCR with GP5+/GP6+ [9, 25, 30, 32, 34, 36] or MY09/MY11 [16, 17, 20] primer set or by nested PCR using primer pairs GP5+/GP6+ and MY09/MY11 [1, 8, 21, 22]. In 7 studies where specific primers for high-risk HPV and low-risk HPV [6, 7, 14, 15, 18, 29, 37] were used, 2 studies used commercial PCR kits (High-risk Human Papillomavirus Multiplex screening genesig® kit and HPV detection Madison® kit) [14, 37], whereas others used specific primers for E6/E7 regions [6, 7, 15, 18, 29]. Also, specific primers for HPV16 and 18 were utilized in one study [7]. Besides, four studies combined detection techniques either by using PCR with immunohistochemistry (IHC) or in situ hybridization (ISH) [19, 33] or by combining ISH with IHC [24, 27].

Generally, HPV16 [1, 7, 21–23, 25, 27, 28, 32, 33, 37] and HPV18 [1, 6, 7, 21, 26, 28, 33] were the most common genotypes detected. Also, cases with HPV33 [1, 20, 26], HPV45 [21, 22], and HPV6 [20, 31] were found. Multiple infections were also reported, in 3 studies a coinfection with two HPV types was reported [21, 22, 29]. Furthermore, in 2 studies, a coinfection with three HPV types was detected [21, 22] (Table 1).

Meta-analysis

In our Meta-analysis, significant heterogeneity between the different included studies was observed: $Q=231.4$, $df(Q)=28$, $I^2=88\%$, $P\leq 0.001$. Q represents the distance of each study from the mean effect, if studies have the same rate of prevalence, Q would be equal to df (degree of freedom). In our study, Q is higher than df , which is an evidence of variation between studies.

Table 1 Subgroup characteristics of HPV in ovarian cancer and borderline lesions by region, HPV genotypes, detection methods, and biological specimens of the included studies

Subgroup	No. of studies	HPV+/total cases	Range of HPV positivity	Prevalence % (95% CI)
Total	29	428/2280	0–81%	15.9 (11–22)
Geographic distribution				
Asia	7	215/839	6–70%	30 (20–44)
Eastern Europe	3	32/134	7.4–81.5%	29.3 (4.4–78)
Middle East	8	175/665	1.9–42%	21.6 (13.2–33.4)
Southern Europe	1	3/71	-	4.225 (1.151–11.702)
Western Europe	4	1/442	0–0.7%	0.226 (0.011–1.270)
North America	6	0/129	-	0 (0–0.02)
Specimen				
FFPE	19	359/1848	0.5%–52%	16.7 (11.6–23.4)
Frozen tissues	8	66/348	0.7–81%	14.1 (3–40)
Fresh tissues	2	4/74	2–6.7%	5.8 (2.3–13.8)
Detection method				
Combined techniques (PCR/IHC; PCR/Southern hybridization; IHC/ISH)	4	50/161	2.6–52%	31.06 (24.4–38.5)
Nested-PCR (L1 Consensus primers GP5+/GP6+ and MY09/MY11)	4	124/426	1.9–42%	29.1 (25–33)
PCR (L1 consensus primer and specific HPV-type primers)	4	91/486	0.7–70%	18.7 (15.5–22.4)
Single-step PCR (L1 Consensus primers GP5+/GP6+ or MY09/MY11)	9	118/819	0.5–35.7%	14.4 (12–17)
PCR (specific primers for HR-HPV and LR-HPV)	7	45/353	0.5–81.5%	12.7 (4.7–10.3)
IHC	1	3/35	-	8.5 (2.9–22.3)
HPV genotypes				
HPV16	15	145/266	0–100%	54 (27.9–55)
HPV18	14	72/304	0–100%	23.6 (18.8–28.26)
HPV33	8	42/346	0–70%	17 (10.2–18.06)
HPV45	7	7/138	0–7.1%	3.8 (0.8–4.3)
HPV6	8	30/246	0–100%	12 (6.9–13.5)
HPV16/18	3	20/116	15.2–100%	17.2 (4.4–20.3)
HPV16/45	2	11/114	2.3–13.8%	9.6 (1.9–10.2)
HPV16/18/45	2	4/114	2.3–4.1%	3.5 (0.3–4.5)
HPV18/45	2	3/114	2.3–2.7%	2.6 (0.2–3.9)

PCR, polymerase chain reaction; HPV, human papillomavirus; FFPE, Formalin-Fixed Paraffin-Embedded; 506 ISH: In-situ hybridization

Because of the substantial heterogeneity found ($I^2 = 88\%$), we applied the random-effect Model, with a pooled rate of 15.9% (95% CI, 11–22). Besides, a large variation of HPV prevalence across the included studies was observed (range 0–81%). The test for the overall effect is $Z = -7.5$, $P \leq 0.001$ (Fig. 2). Thus, the null hypothesis was rejected that the true prevalence of HPV is comparable between the included studies.

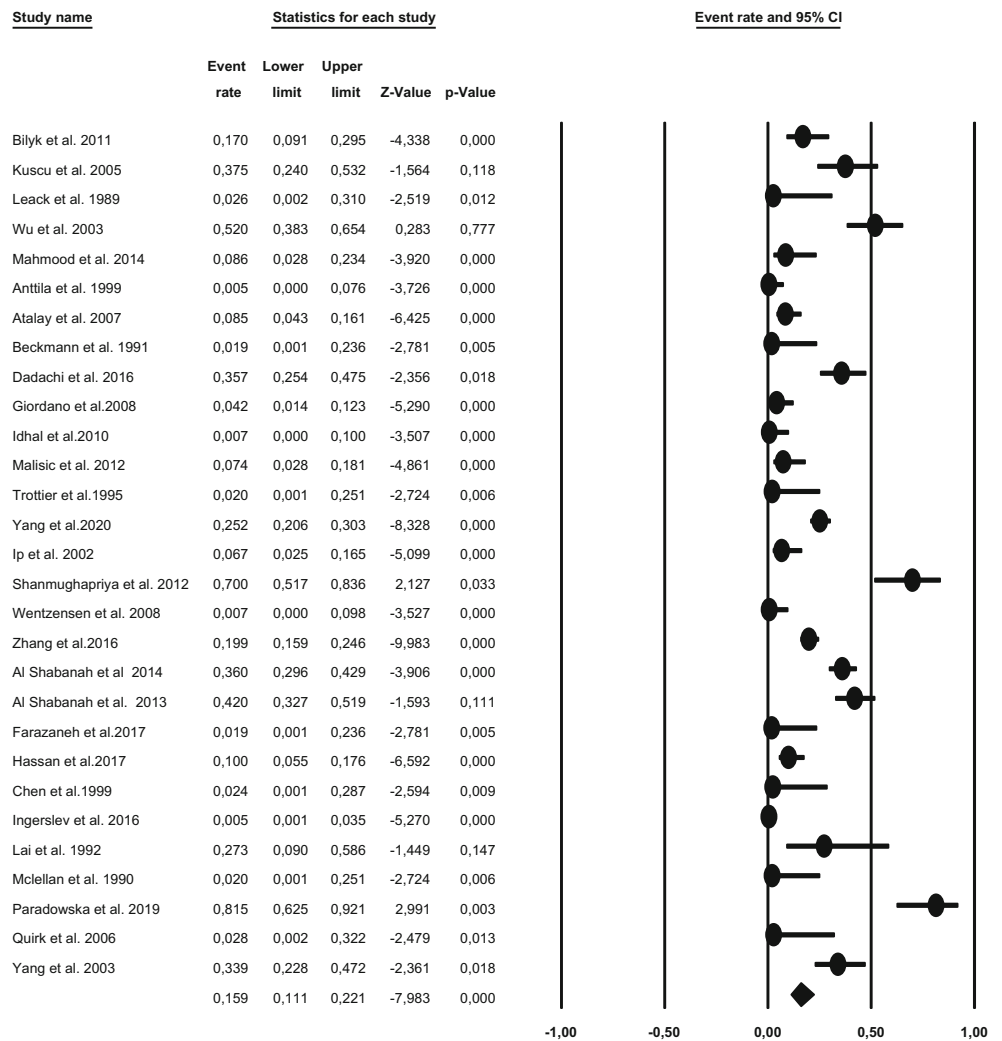
Subgroup and meta-regression analyses

A subgroup analysis was performed to investigate if the results found for all studies included would apply separately for the stratified data according to geographical region, type of

specimen, or detection method. The results of the subgroup analysis are shown in Tables 1 and 2.

Stratifying by geographical region, the highest rates of HPV detection in ovarian cancer tissue were found in Asia (30.9%; 95% CI, 20–44) ranging from 6% to 70%, Eastern Europe (29.3%; 95% CI, 4.4–78; range, 7.4–81.5%), and in the Middle East (21.6%; 95% CI, 13–33.4; range, 1.9–42%). The lowest prevalence was detected in North America (0%; 95% CI, 0–0.02) (Table 1). Also, the test of the null hypothesis yields a Q -value = 59.4, $df(Q) = 5$, and P -value < 0.001 (Table 2), indicating that the effect size differs by geographical regions. Of note also within a given region the variation in prevalence is substantial.

Fig. 2 Forest plot showing HPV rate and 95% confidence intervals in ovarian cancer studies based on random-effect model



Results of subgroup analysis based on specimen types suggest that HPV rate is slightly greater when FFPE tissues were examined (16.7%; 95% CI, 11.6–23.4) compared to frozen tissue (14.1%; 95% CI, 3–39.9) or fresh tissue (5.8%; 95% CI, 2.3–13.8). The overall *Q*-value, *df(Q)* and *P*-value for specimen type subgroup were 4.9 with 2 *df* and 0.08, respectively. Thus, statistically, the effect size does not differ by the specimen type.

Regarding detection methods of HPV, the lowest prevalence of HPV is found when IHC is used (8.5%; 95% CI, 2.9–22.3) [23]. A greater prevalence was observed when more sensitive or combined techniques (PCR and ISH or IHC; southern hybridization and IHC) (31%; 95%CI, 24.4–38.5) were utilized. For the detection method stratification, the overall *Q*-value is 31 with 5 *df*, and the corresponding *P*-value is 0.1. Thus, the effect size does not differ significantly by the detection method.

Taken together, the meta-regression analyses showed a significant association of HPV prevalence with geographic distribution ($P \leq 0.001$). However, the specimen type ($P=0.19$)

and HPV detection methods ($P=0.3$) were not statistically associated with HPV prevalence.

Sensitivity analysis

To detect a potential bias related to the quality of the included studies, we performed a sensitivity analysis by calculating HPV prevalence when excluding one study at a time. Figure 3 shows that after the exclusion of each study, no significant effect on the overall prevalence of HPV in ovarian cancer patients was observed.

Publication bias

A visual inspection of the funnel plot showed an asymmetry, and a one-tailed *p*-value of Begg’s test and Egger’s regression test were 0.006 and 0.5, respectively (Figure 4). Suggesting that there is a significant publication bias.

Table 2 Subgroup analysis by categorizing the data according to the geographic region, HPV genotypes, detection methods, and biological specimens of the included studies

Subgroups	Test of null (2-tail)		Heterogeneity		
	Z value	P-value	Q-value	df(Q)	P-value
Geographical region					
Asia	-2.7	0.006			
Eastern Europe	-0.7	0.4			
Middle East	-4.2	<0.001			
North America	-6.460	<0.001			
Southern Europe	-5.290	<0.001			
Western Europe	-8.145	<0.001			
Overall	-10.123	<0.001	59.4	5	<0.001
Specimen					
FFPE	-7.4	<0.001			
Fresh Tissues	-5.7	<0.001			
Frozen tissues	-2.5	0.01			
Overall	-9.1	<0.001	4.9	2	0.08
Detection method					
PCR (L1 Consensus primers GP5+/GP6+ and MY09/MY11)	-2.7	0.007			
PCR (L1 Consensus primers GP5+/GP6+ or MY09/MY11)	-5.9	<0.001			
PCR (specific primers for HR-HPV and LR-HPV)	-2.5	0.01			
PCR (L1 consensus primer and specific HPV-type primers)	-9.9	0.04			
Combined techniques (PCR/IHC; PCR/Southern hybridization; IHC/ISH)	-1.8	0.06			
IHC	-3.92	<0.001			
Overall	-7.9	<0.001	31.8	5	0.1

PCR, polymerase chain reaction; HPV, human papillomavirus; FFPE, Formalin-Fixed Paraffin-Embedded; 506 ISH: In-situ hybridization

Discussion

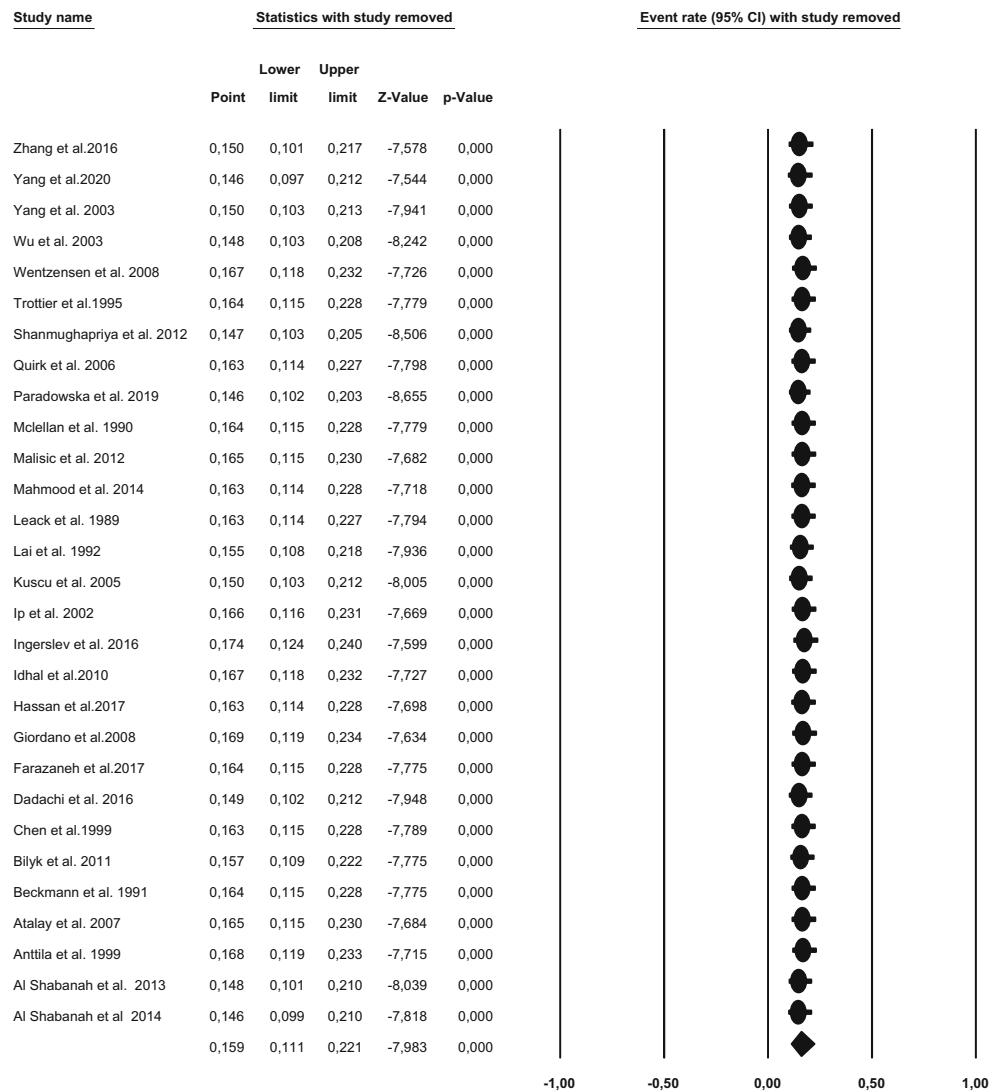
Since the first report on HPV infection in ovarian cancer published by Kaufman et al. in 1987 [38] using southern blot hybridization, researchers have explored the potential involvement of HPV in ovarian cancerogenesis in several geographical regions and by more sensitive and specific methods. However, the association between HPV and ovarian cancer remains controversial. In the current meta-analysis, 2280 cases with ovarian cancer were investigated. In 19/29 studies where HPV was detected, 2 reported a cause-effect relation by detecting E6/E7 oncogenes expression or HR-HPV viral integration [21, 31]; while in 10 studies no HPV was found. The overall pooled prevalence of HPV was 15.9% (95% CI, 11–22), and the prevalence of detection ranged from 0 to 81%. This result is in accordance with previous meta-analyses published in 2013 [2, 10]. Nevertheless, in the present meta-analysis, six publications from the Middle East and four recent studies were additionally included.

This overall low prevalence of HPV found in ovarian cancer compared to the strongly HPV-associated malignancies, i.e., cervical cancer, might be explained by the adaptation and affinity of HPV to certain squamous cells in the cervical epithelium, or/and by the vulnerability of the transformation zone of the cervix. The latter is a niche of cells with an embryonic characteristic such as cuboidal epithelial cells, reserve cells, or potentially embryonal stem cells, which have been proposed as targeted cells by HPV infection [39–43]. However, HPV target cells in the ovarian epithelium have not been described yet.

Substantial heterogeneity between studies was found ($I^2=88\%$, $P \leq 0.001$) and an asymmetric funnel plot was observed. These findings suggest a wide variation between study results. This is supported by major differences in terms of sample size, geographic distribution, type of specimen, detection methods, and the period of the study. Thus, to balance the selection bias, subgroup analyses were carried out.

It has been reported that geographical differences in the prevalence of HPV might be due to biological and geographic

Fig. 3 Sensitivity analysis showing the impact of exclusion of any study on the summary effect



interaction between HPV variants and the host immunogenetic factors such as GST (glutathione-S-transferase), HLA (human leukocyte antigen), MDM2 (Mouse double minute 2 homolog), FAS (fatty acid synthase) gene promoter-670 and p53 codon 72 polymorphisms [44, 45]. Furthermore, impairment of the cellular immunity through malnutrition, inflammation, HIV, and microbial infection, genetic predisposition, lifestyles; might contribute to a higher incidence of HPV acquisition in some regions [46]. Moreover, in cervical cancer, an association between the prevalence of high risk-HPV infection in the general population and the incidence of this malignancy has been demonstrated [47]. Also, several researchers have found regional differences in the prevalence of HPV genotypes in invasive cervical cancer specimens. In a meta-analysis conducted on 85 studies, HPV16 was predominant in cervical squamous cell carcinoma, with a prevalence of 63% in North America and 46% in Asia [47]. Similarly, in a study conducted on 10,575 cases of invasive cervical cancer from 38 different countries tested for HPV, researchers found that

HPV16 and 18 had the highest relative contributions in North America (79%) and Oceania (79%), followed by Africa (71%) and Asia (71%) [48]. In our subgroup analysis, stratified by geographical distribution, studies carried out in Asia (30.9%; 95% CI, 20–44) and Eastern Europe (29.3%; 95% CI, 4.4–78) had the highest prevalence of HPV in ovarian cancer detected while no HPV association was detected in North America. Furthermore, the meta-regression showed that geographic distribution explains the variation in worldwide HPV prevalence described in ovarian cancer.

FFPE blocks are a valuable source to conduct retrospective studies. However, a prolonged fixation time with formalin causes a cross-linking of proteins and nucleic acids and also random breakages in nucleotide sequences [49]. Besides, fixation protocols including non-buffered formalin may contribute to decreasing DNA integrity and reduced HPV detection rate. Furthermore, false positivity due to carryover contamination might occur if FFPE blocks were not processed under strict conditions [48]. On the other hand, frozen tissues allow

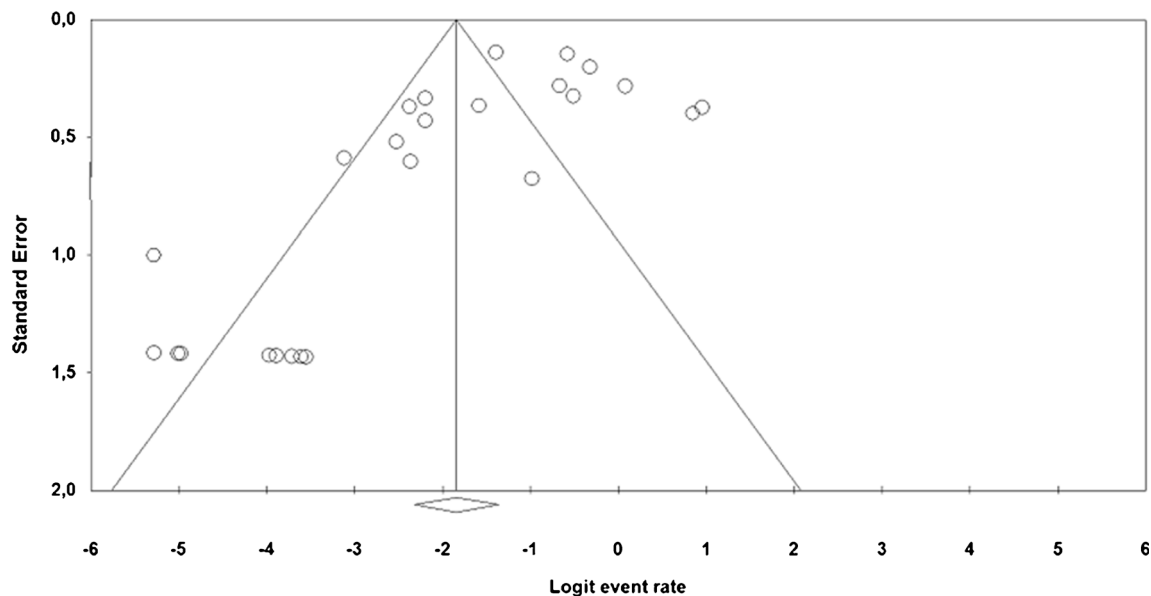


Fig. 4 Funnel plot of studies included in the meta-analysis. The distribution of the studies (dots) is asymmetric suggesting an evidence of publication bias that is also found by Egger's regression test ($P=0.006$) and Begg and Mazumdar rank correlation ($P=0.5$)

better DNA preservation, but require constant low-temperature maintenance [50]. In the present study, the HPV detection rate was slightly higher when frozen tissues were used. Nevertheless, the subgroup and the meta-regression analyses showed that specimen type does not explain the variation of HPV prevalence.

Currently, HPV detection relies on molecular techniques that allow a high sensitivity and specificity of detection [51]. Usually, HPV detection by PCR can be performed either by using a single PCR primer set such as (i) L1 consensus primers MY09/MY1, GP5+/GP6+, or SPF10 primers followed by type-specific probe-based detection, direct sequencing or restriction-fragment length polymorphism (RFLP), (ii) by nested PCR using a combination of MY/GP primers; (iii) or by using type-specific primers [52]. Nevertheless, a single PCR primer set underestimates the prevalence of HPV positivity, but the combination of MY/GP in nested PCR is considered the most sensitive DNA-based test that allows detection of low levels of HPV [53]. In the present meta-analysis, PCR was the most often used technique (in 26/29 studies). However, the prevalence of HPV in ovarian tissues was the highest when 2 different techniques were combined (PCR and ISH or IHC) (31%; 95%CI, 24.4–38.5), and when nested PCR GP5+/GP6+ followed by reamplification with MY09/MY11 was used (29.1%; 95%CI, 25–33). Furthermore, a lower prevalence of HPV was found by using type-specific primers either by targeting E6/E7 regions or by using designed type-specific primers (12.7%; 95%CI, 4.7–10.3).

The overexpression of E6 and E7 oncoproteins are required to maintain malignant transformation in HPV-related cancers, which should be detected by specific primers for E6/E7 regions [54, 55]. Contrary to the well-established pathway of E6

and E7 in the cervical cancerogenesis, it has been discussed that the molecular mechanism of cancerogenesis in HPV-infected ovarian epithelium operates differently. This hypothesis was stated by Kisseljova et al. [53], who were unable to detect HPV16 DNA by E6/E7 primers despite finding HPV-association in their samples investigated. Also, Roos et al. [56] reported evidence of HPV presence in ovarian cancer in North America by detecting transcripts of HPV18 oncogenes in ovarian cancer transcriptomes (by RNA-Seq) from The Cancer Genome Atlas (TCGA) database. Hence, a more sensitive assay, based on mRNA HPV oncogene expression is needed, to clarify the role of this microorganism in ovarian cancerogenesis by comparing viral oncogene expression in cancer samples and adjacent normal tissues.

According to Bosch et al. [57], HPV16 is the most frequent genotypes detected in cervical cancer cases in most countries (50% to 60%), followed by HPV18 (10–20%), HPV45 (4–8%), and HPV31 (1–5%). Comparable results were found in our subgroup analysis stratified by HPV genotypes.

Coinfection with more than one HPV anogenital type was observed in more than 50% of sexually active women through their life [58]. Several studies showed that the presence of multiple HPV genotypes is associated with an increased risk of high-grade lesions at the cervix [59–61]. Besides, coinfections with alpha-9 genotypes increase the risk of cervical cancer 5.3 fold compared to a coinfection with alpha-7 genotypes [62]. Moreover, Senapati et al. [63], revealed that women infected with different genotypes in a phylogenetically related clad had a higher risk of cervical carcinoma in comparison to women infected with unrelated phylogenetic clad. Also, an in vitro study reported that coinfection of a single cell with HPV16 and HPV18 induces a replication interference

between them [64]. Our meta-analysis revealed that an HPV16/18 coinfection was the most observed (17.2%) followed by HPV16/45 (9.6%). However, in one study where a coinfection with HPV16/18 was reported [29], 2 of 3 positive cases had an HPV16/18 coinfection that could potentially be a result of cross-contamination while sectioning FFPE blocks, knowing that no blank FFPE block was sectioned between cases to reveal any carryover contamination. Indeed, in order to prove an ethiological involvement of HPV in a certain cancer development, e.g. in HNSCC, recently a very elaborate and strict methodological process has been described involving (i) control for carryover by HPV free tissue sections between experimental blocks, (ii) measurement of HPV oncogene expression by detection of E6 and/or E7 mRNA, and (iii) detection p16^{ink4a} upregulation [65].

There were several limitations in our meta-analysis. First, in 22 studies, no control group was selected and the sample size was small in other studies, which does not allow to distinguish the difference between malignant and control groups. Second, few studies, reported HPV genotypes by histological type. The latter points out the need for more studies that take into account the histologic type of the malignancy. Third, no HPV oncogene expression was measured to investigate its role in ovarian cancerogenesis by comparing viral oncogene expression in malignant and normal tissues. Fourth, no carry-over contamination control was performed while processing FFPE blocks, which might lead to an overestimation of HPV prevalence. Fifth, coinfection with more than one HPV genotype was evaluated in just 3 studies which did not allow us to investigate if the presence of multiple HPV genotypes could be associated with an increased risk of ovarian cancer. Sixth, the heterogeneity between studies was substantial and publication bias was found, this might be explained by methodological differences and the fact that scientific journals are more likely to publish studies that report positive results. Seventh, the majority of the included studies reported the studied population age in mean, so we were unable to perform a subgroup analysis by age group or perform meta-regression analysis, thus the discrepancy in prevalence could be due to differences in the age range of included studies. Furthermore, the data discrepancy might be also explained by the variation of HPV genotypes across geographical regions. Eighth, none of the included studies has mentioned if patients received HPV vaccine, which could be of interest to assess the efficacy of prophylactic vaccines for the potential prevention of HPV infections in the upper female genital tract.

Considering that HPV plays a role in a rising number of head and neck squamous cell carcinomas, it has been reported that patients with HPV+ tumors have a favorable prognosis, and HPV-selective therapies are under investigation. If there is a potential implication of HPV in ovarian cancerogenesis, treatment and maybe prognosis, HPV-association of ovarian cancer needs further investigation.

Conclusion

Our meta-analysis suggests a great difference in the prevalence of HPV detected in ovarian cancer which is not seen in strongly HPV-associated cancers such as cervical cancer. However, further studies are needed, using more precise assays that identify active infection by testing HPV oncoprotein expression and informative biomarkers like p16 upregulation to prove the causality of HPV detection with cellular transformation. Also, other covariates such as ethnicity, age, and lifestyle have to be considered.

Supplementary Information The online version contains supplementary material available at <https://doi.org/10.1007/s10096-021-04282-7>.

Availability of data and materials The datasets and referenced literature supporting the conclusions of this article are included within the article and its additional supplementary file.

Author contribution Conception and design of the study: S.C., A.M.K., and J.S.; methodology: S.C., A.M.K., A.A., S.T., and E.T.; formal analysis: S.C.; Search: S.C. and A.A.; data extraction: S.C. and A.A. writing: S.C.; review and editing: A.M.K.; supervision: A.M.K., J.S., and E.B. All author(s) read and approved the final manuscript.

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Declarations

Ethics approval Not applicable

Consent for publication Not applicable

Conflict of interest The authors declare no competing interests.

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