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A practical approach to the sensitivity analysis for kinetic Monte Carlo simulation of heterogeneous catalysis

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Lattice kinetic Monte Carlo simulations have become a vital tool for predictive quality atomistic understanding of complex surface chemical reaction kinetics over a wide range of reaction conditions. In order to expand their practical value in terms of giving guidelines for the atomic level design of catalytic systems, it is very desirable to readily evaluate a sensitivity analysis for a given model. The result of such a sensitivity analysis quantitatively expresses the dependency of the turnover frequency, being the main output variable, on the rate constants entering the model. In the past, the application of sensitivity analysis, such as degree of rate control, has been hampered by its exuberant computational effort required to accurately sample numerical derivatives of a property that is obtained from a stochastic simulation method. In this study, we present an efficient and robust three-stage approach that is capable of reliably evaluating the sensitivity measures for stiff microkinetic models as we demonstrate using the CO oxidation on RuO₂(110) as a prototypical reaction. In the first step, we utilize the Fisher information matrix for filtering out elementary processes which only yield negligible sensitivity. Then we employ an estimator based on the linear response theory for calculating the sensitivity measure for non-critical conditions which covers the majority of cases. Finally, we adapt a method for sampling coupled finite differences for evaluating the sensitivity measure for lattice based models. This allows for an efficient evaluation even in critical regions near a second order phase transition that are hitherto difficult to control. The combined approach leads to significant computational savings over straightforward numerical derivatives and should aid in accelerating the nano-scale design of heterogeneous catalysts. *Published by AIP Publishing.* [<http://dx.doi.org/10.1063/1.4974261>]

I. INTRODUCTION

The last few years have seen tremendous progress in modeling and analyzing heterogeneous catalysis using the first-principles kinetic Monte Carlo (1p-kMC) approach.¹ The appealing features of the approach are an elementary reaction mechanism (and corresponding rate constants), which has been derived from predictive quality electronic structure methods, and a subsequent solution of the resulting master equation by trajectory sampling. The last step thereby only introduces a numerical (tunable) error, in contrast to the prevalent phenomenological kinetics where the employed mean-field approximation might miss qualitative features.^{2,3} However, kinetic Monte Carlo has some significant drawbacks. Besides the need to perform one reaction event after the other and the concomitant large computational costs, it is a huge effort to obtain parameter sensitivities,

which are a measure for the rate-determining steps.^{4,5} The evaluation is generally costly because sensitivities are not straightforward expectation values such as coverages or rates.

During the last few years, a number of approaches for the sensitivity analysis of kMC models have been developed, mostly in the context of biological “reaction” networks and the chemical master equation (in this context kMC is often termed as stochastic simulation). These can be grouped into two sets of approaches: (i) targeting at a reduction of the noise in finite difference approximations^{6–9} and (ii) targeting a direct estimation of sensitivities from the analysis of the simulated trajectories.^{10–12} Albeit being improvements to the brute force numerical differentiation, both groups have their limitations. Direct estimation approaches allow obtaining all sensitivities from the same set of trajectories (for ergodic stationary processes from a single trajectory). On the other hand, the variance of the corresponding estimators increases with increasing time horizon.^{6,11} Especially for stiff problems, the latter point is a major concern, as it severely affects the estimation of steady state sensitivities, and there are quite some efforts on sampling strategies for stationary sensitivities.^{11,12} This problem is not as severe for the finite difference based approaches. These, however, require additional simulations for every targeted

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sensitivity and generally a good guess for the difference parameter.¹³

For lattice based models as they appear in heterogeneous catalysis, the methodology for sensitivity analysis is not that mature and there are only a limited number of studies addressing these high-dimensional, usually very stiff problems.^{7,14–17}

In this study, we devise a three-stage strategy for the estimation of stationary sensitivities, which is suitable for these problems. We illustrate this strategy on the model for the CO oxidation on RuO₂(110) by Reuter,¹⁸ which is a popular fruit-fly test problem.^{19,20} We revisit the reaction conditions studied in Ref. 16 and can therefore concentrate on discussing the peculiarities of the sensitivity estimation.

The basic idea of the approach is to first try to directly sample the sensitivities from a single trajectory and to only employ the Coupled Finite Differences (CFDs) method⁶ for those sensitivities, for which the direct sampling estimator shows a too high variance. In practice, directly sampling all sensitivities without prior knowledge might soon become unpractical. We therefore first estimate bounds for the sensitivities using the pathwise relative entropy method¹⁷ (Section III A). This also allows to extract an expansion parameter, which is needed for our direct sampling approach and represents the system's memory. So, we avoid its estimation in the direct sampling step and can significantly reduce the sampling effort there. Our direct sampling approach is based on a series expansion of integrated linear response functions and is presented in Section III B. As it turns out, the derived estimator is very similar to that by Chen and Cao.^{10,21} However, our estimator is based on a very different reasoning and avoids the use of random time steps. In the end, these two steps allow to significantly reduce the number of required Coupled Finite Difference (CFD) estimates as presented in Section III C.

With the addition of the direct sampling, our approach can be viewed as an extension of the approach put forward by Arampatzis, Katsoulakis, and Pantazis in the context of the chemical master equation.²² We find this additional step to be beneficial, as, for most of the considered reaction conditions, none or just very few CFD estimates remained necessary.

II. BACKGROUND

Sensitivity analysis aims at identifying the most important input parameters for a computational model. That is, if we change an input parameter, how much does the model output change? This information can then be used to identify those parameters, which need to be more accurately determined in order to arrive at a more reliable model. Or, we can employ it to find out, which aspects of the underlying physical system could be optimized in order to arrive at a better performance. In the first order, such sensitivity information is provided by partial derivatives of the model output with respect to the input parameters.

In the microkinetic modelling of heterogeneous catalysis, the central model outputs are the (stationary) reaction rates of one or more target reactions. As the input parameters have an atomistic meaning, sensitivity analysis is a tool to determine

which atomistic aspects of the catalyst control the macroscopic reactivity.

In this section, we outline the background for performing sensitivity analysis for 1p-kMC models. We start with the Markov jump process description of kinetics on lattices. We then introduce sensitivity analysis in this context. After introducing the kinetic Monte Carlo methodology as a numerical tool for simulating Markov jump processes, we detail our test problem, the CO oxidation on RuO₂(110).

A. Master equation

In chemical kinetics, the rare event dynamics, typical for surface catalytic processes, is exploited by considering a time evolution that is coarse-grained to the discrete elementary processes of the reaction mechanism. In this description, a state j of the system corresponds to a meta-stable state or domain in the microscopic evolution. That is, the time which is spent between two transitions between these domains is large compared to the time spent for one transition. We can then safely assume that in the time between such rare events the system loses all memories of the past by thermal fluctuations and the sequence of states can then be regarded as a Markov jump process. The probabilities P_i of finding the system in the state i then obey a master equation,^{23,24}

$$\frac{d}{dt}P(t) = \Gamma P(t), \quad (1)$$

where $P(t) = (P_1(t), P_2(t), \dots)^T$ is the vector of the probabilities. The stochastic generator Γ has the matrix elements,

$$\Gamma_{ij} = w_{ij} - \delta_{ij} \sum_l w_{li}, \quad (2)$$

where w_{ij} is the transition rate for the event $j \rightarrow i$, with $w_{ii} = 0$. The (negative) diagonal element $w_i^{\text{acc.}} = \sum_l w_{li}$ is called the accumulated rate (for the state i), which simply is the rate for escaping the state i . In the following, we consider processes, which relax to a single stationary distribution $P_{\text{stat.}}$. Further, we assume that the transition probability $P_{\text{kMC}}(j|i) = w_{ji}/w_i^{\text{acc.}}$ defines a discrete time Markov chain, which also relaxes to a single stationary distribution. For these prerequisites, we can prove the convergence of the series expansion employed in Sections III A and III B (for details see the [supplementary material](#)).

In the context of catalysis, the state i can be regarded as an integer vector, which carries the information that which species is adsorbed on which adsorption site at the surface. Due to symmetries, like translational invariance, and the locality of processes, we can group the allowed events $i \rightarrow j$ ($w_{ij} > 0$) into different subsets, each assigned to one elementary step (or reaction). This decomposition can be made a partition of the set of all allowed events, i.e., every allowed event is in a single set α . For convenience, we impose that all processes $j \rightarrow i$ belonging to the same subset α have the same value for their transition rate. Thus for each reaction α , we can define the partial transition matrix w_{ij}^α with the elements

$$w_{ij}^\alpha = \begin{cases} k^\alpha, & \text{if } j \rightarrow i \in \alpha, \\ 0, & \text{otherwise,} \end{cases} \quad (3)$$

where k^α will be called the rate constant (RC) of the reaction α in the following. With w_{ij}^α we can define the (partial) generator Γ^α as

$$\Gamma_{ij}^\alpha = w_{ij}^\alpha - \delta_{ij} \sum_l w_{li}^\alpha. \quad (4)$$

Summing over all (partial) generators, we arrive at the (total) generator Γ ,

$$\Gamma_{ij} = \sum_{\alpha=1}^{N_{\text{reac.}}} \Gamma_{ij}^\alpha, \quad (5)$$

where $N_{\text{reac.}}$ is the total number of reactions we have for our system. As for the total generator, the absolutes of diagonal elements of the partial generator are the partial accumulated rates $w_i^{\alpha, \text{acc.}} = \sum_l w_{li}^\alpha$.

B. Sensitivity analysis

The central objective of kinetics in the heterogeneous catalysis is average, stationary reaction rates $\langle R \rangle$, also called turnover frequencies (TOFs), if suitably normalized. In the case that there are multiple elementary reactions contributing to the targeted overall reaction, $\langle R \rangle$ is the superposition of the corresponding reaction rates $\langle R^\alpha \rangle$ of the elementary reactions,

$$\langle R \rangle = \sum_{\alpha}^{N_{\text{reac.}}} T^\alpha \langle R^\alpha \rangle, \quad \text{with } \langle R^\alpha \rangle = \sum_i w_i^{\alpha, \text{acc.}} P_{i, \text{stat.}}, \quad (6)$$

where $P_{i, \text{stat.}}$ is the stationary probability distribution and T^α is a constant which only depends on the stoichiometry of the reaction network. For the below discussed CO oxidation on RuO₂(110), there are four reactions which yield one CO₂ molecule per event. In this case, T^α for the CO oxidation rate is one for each of these four and zero for all other reactions.

We require a unique stationary probability distribution $P_{\text{stat.}}$ and the reaction rate $\langle R \rangle$ is therefore a function of the rate constants. As the RCs measure how fast the elementary reactions proceed, finding out how $\langle R \rangle$ reacts on little changes in the RCs can be used to identify the most important steps, i.e., those steps which need to be accelerated or slowed down to achieve a higher reaction rate. In chemical kinetics, a useful measure for such sensitivity analysis is¹⁶

$$X^\alpha := \frac{k^\alpha}{\langle R \rangle} \left(\frac{\partial \langle R \rangle}{\partial k^\alpha} \right)_{k^{\beta \neq \alpha}}, \quad (7)$$

i.e., the relative change of $\langle R \rangle$ over the relative change of the corresponding RC, while all other RCs are kept constant. We will term this sensitivity index Degree of Rate Sensitivity (DRS).

The DRS can be regarded as the generalized reaction order, which becomes clear for the case that all DRS do not change for a range of values for the RCs. In this range, the reaction rate obeys the relation

$$\langle R \rangle = r_0 \prod_{\alpha}^{N_{\text{reac.}}} (k^\alpha)^{X^\alpha}, \quad (8)$$

where r_0 is independent of the RCs. For elementary steps involving gas phase species A like adsorption or Eley-Rideal

reactions, we expect rate expressions of the form $k^\alpha \propto p_A$ with the partial pressure p_A . We then arrive at a power law kinetics

$$\langle R \rangle = \tilde{r}_0 \prod_A (p_A)^{\nu_A}, \quad \text{with } \nu_A = \sum_{\alpha \in R_A} X^\alpha, \quad (9)$$

where R_A is the set of elementary steps involving the gas phase species A and \tilde{r}_0 is independent of the partial pressures. Thus the DRS connects macroscopic reaction orders with microscopic elementary steps and rate constants. Usually, we employ the rate expressions $k^\alpha = f^\alpha(T, \{p_A\}) \exp(-\Delta E^\alpha / k_B T)$, where k_B is the Boltzmann constant, T is the temperature, and ΔE^α is the activation barrier for the reaction. Using sensitivity analysis, we can find a simple connection between the elementary barriers and the macroscopic apparent activation barrier E_{app} by¹⁶

$$E_{\text{app.}} := - \left(\frac{\partial \langle R \rangle(T, \{p_A\})}{\partial (k_B T)^{-1}} \right)_{p_A} \approx \sum_{\alpha}^{N_{\text{reac.}}} X^\alpha \Delta E^\alpha. \quad (10)$$

For brevity, we have only provided the dominant contribution and neglected the (typically weak) temperature dependence of the pre-exponential factor.

To simplify the latter discussion, we decompose the DRS,

$$\frac{k_i}{\langle R \rangle} \left(\frac{\partial \langle R \rangle}{\partial k^\alpha} \right)_{k^{\beta \neq \alpha}} = X_0^\alpha + X_1^\alpha \quad (11)$$

using the product rule. Here $X_0^\alpha = T^\alpha \langle R^\alpha \rangle / \langle R \rangle$, which is the relative contribution to the total rate of the reaction α . As this is a stationary expectation, it is straightforward to obtain from sampling. Therefore we will concentrate on the second term,

$$X_1^\alpha = \frac{k_i}{\langle R \rangle} \sum_i R_i \frac{\partial P_{i, \text{stat.}}}{\partial k^\alpha} \quad (12)$$

in the further discussion. Intuitively, X_1^α can be regarded as the average effect on the creation and annihilation of states i , which allow for those reactions contributing to the overall reactivity. As X_1^α originates from the k^α -dependence of the probability distribution, this cannot be expressed as a stationary expectation, except for the case when the stationary distribution is known as a function of the RCs. For obtaining X_1^α , we can omit any explicit dependence of R on variations of the RCs.

The DRS is strongly related to the Campbell's Degree of Rate Control (DRC),⁴

$$X_{\text{DRC}}^\xi = \frac{k_+^\xi}{\langle R \rangle} \left(\frac{\partial \langle R \rangle}{\partial k_+^\xi} \right)_{k_+^{\phi \neq \xi}, K^\phi}, \quad (13)$$

where ξ denotes a pair of forward and backward reaction and k_+^ξ and K^ξ are the corresponding RCs of the forward reaction and the equilibrium constant, respectively. The only difference between DRC and DRS is what is kept constant during differentiation, all other RCs (DRS) or all other forward RCs and the equilibrium constants (DRC). In a previous publication, we therefore termed both the degree of rate control.¹⁶ To clarify the presentation, we have introduced a new name for the less common sensitivity measure (7). Applying the chain rule, we arrive at a simple relation between DRS X^α and DRC X_{DRC}^ξ ,¹⁶

$$X_{\text{DRC}}^\xi = X^{\xi+} + X^{\xi-}, \quad (14)$$

where ξ_+ identifies the forward reaction and ξ_- the backward reaction. The DRC X_{DRC}^{ξ} is therefore easily obtained by adding the DRSs for the forward and the reverse reactions.

By construction, the DRC X_{Camp}^{ξ} is zero for reactions which are in equilibrium. Identifying rate-determining steps by large $|X_{\text{DRC}}^{\xi}|$ agrees then with chemical intuition, which would not assign the property “rate-determining” to steps which are in equilibrium. Also the DRC has the nice microscopic interpretation of the effect of a change in the height of the barrier separating two metastable states on a potential energy surface. This microscopic interpretation in terms of barriers and the thermodynamic consistency are absent in definition 7 for X^{α} . On the other hand, there are many situations, which cannot be properly addressed using the DRC. If we change the partial pressures of the gas phase species above the catalyst, adsorption RCs are affected but not desorption RCs and thus the equilibrium constants change. If we modify the catalyst or we exchange the electronic structure method for estimating the RCs, we do not only change the barrier height but also the equilibrium properties. As long these changes do affect the mechanism, they are reflected in the Markov model by the values of the RCs and the DRSs X^{α} provide the sensitivity to all rate constants. Using the chain rule, the DRS therefore allows to calculate any sensitivity measure, which is adapted to the problem at hand. In this context, we have already provided explicit formulas for three special cases: (i) pressure variation (reaction orders, Eq. (8)), (ii) temperature variation (apparent activation barrier, Eq. (10)) and (iii) barrier height variation (DRC, Eq. (14)). So, the DRS is much more general allowing to calculate the DRC as a special case. Especially, the kinetic key properties, reaction order 8 and apparent activation energy 10 cannot be extracted from the DRC, as they correspond to changes in the equilibrium constants. Despite its generality, the DRS also provides a decomposition of the DRC into contributions from the forward and the backward RCs by Eq. (14). This decomposition often allows to explain why a step is rate determining and to identify additional key aspects controlling reactivity.

Because of its generality, we will concentrate on the DRS for most of the manuscript, also for the reason of shorter equations and a more comprehensive and compact presentation of the theoretical background. However, most information can probably be extracted using both, the DRC for the identification of the rate-determining steps and the DRS to gain a deeper insight into the details of the mechanism. We will therefore come back to the DRC in Section V when we discuss our final results in terms of surface coverage and barriers. By Equation (14), this comes at essentially zero computational costs.

C. Kinetic Monte Carlo

In most cases, the master equation is too high-dimensional for a direct numerical solution. Instead, one will simulate the underlying Markov process and obtain estimates of the targeted expectations by averaging over trajectories. For this and all our implementations, we use the kmos code package,²⁵ which we have developed during the last years for the lattice based kinetic Monte Carlo.

We employ the so-called variable-step size method,²⁴ which can be described as follows:

1. Initialize the state i and the time t .
2. Increase the time $t \rightarrow t + \delta t$ with $\delta t = (w_i^{\text{acc.}})^{-1} \log \xi$, where ξ is a uniformly distributed random variable on (0,1].
3. Choose the event $i \rightarrow j$ with the probability $P_{\text{kMC}}(j|i) = w_{ji}/w_i^{\text{acc.}}$.
4. Set $i = j$ and go to 1 or stop if the termination criterion is reached.

The method essentially simulates a classical discrete time Markov chain with the transition probability matrix P_{kMC} and simply adjusts the time steps between the jumps to arrive at a statistically proper continuous time Markov jump process.

Focusing on stationary reaction rates, we can most conveniently obtain estimates by time averaging. That is, we will approximate $\langle R^{\alpha} \rangle$, after a sufficiently long relaxation, with

$$\langle R^{\alpha} \rangle \approx \frac{1}{\tau_N} \sum_{n=0}^N w_{i_n}^{\alpha, \text{acc.}} \delta t_n \approx \frac{1}{t_N} \sum_{n=0}^N w_{i_n}^{\alpha, \text{acc.}} \Delta t_n, \quad (15)$$

where (i_1, \dots, i_N) is the sequence of states generated by the kMC simulation and $(\delta t_1, \dots, t_N)$ are the corresponding time steps with $\tau_N = \sum_{n=0}^N \delta t_n$. For the second approximation, we have pre-averaged the time steps, i.e., $\Delta t_n = (w_{i_n}^{\text{acc.}})^{-1}$ and $t_N = \sum_{n=0}^N \Delta t_n$. This is a common approximation^{11,12} to reduce the noise caused by the random time steps and the introduced bias vanishes for large enough N .

In practice, we will employ a number independent sample trajectories to estimate an expectation $\langle A \rangle \approx \bar{A} = K^{-1} \sum_k A_k$ and the standard deviation of our estimate $\sigma(\bar{A})^2 \approx (K^2 - K)^{-1} \sum_k (A_k - \bar{A})^2$. Here K is the number of samples and A_k are k time averages. We generate these samples by first simulating one trajectory. We then perform a number of kMC steps for decorrelation and then use the last configuration as the initial condition for the next trajectory. We repeat this until we have reached the total number K of samples.

D. CO oxidation on RuO₂(110)

Throughout this article, we consider the realistic first-principles kinetic Monte Carlo (1p-kMC) model for the CO oxidation on RuO₂(110) by Reuter.¹⁸ It is one of the most thoroughly tested and understood 1p-kMC models (e.g., Refs. 3, 19, and 26) for which we have already performed a sensitivity analysis.¹⁶ This makes it well suited for testing our scheme and we can concentrate on discussing the peculiarities of the employed sampling approaches.

In detail, the model considers two kinds of adsorption sites, the so-called bridge (br) and coordinatively un-saturated (cus) sites. These are arranged on a square lattice with alternating rows. Oxygen and CO can adsorb on both kinds of sites, where oxygen adsorbs dissociatively while CO stays intact after adsorption. CO₂ does not bind to the surface and the sites can be in three different states: (i) empty, (ii) CO occu-

pied, and (iii) atomic oxygen occupied. Altogether, the model comprises 22 elementary reaction steps, which are single site CO ad/desorptions, nearest-neighbors O₂ ad/desorptions, adsorbed CO/O nearest-neighbor diffusion, and formation of gaseous CO₂ by Langmuir-Hinshelwood reactions of co-adsorbed CO and O. The corresponding RCs have been obtained from the density functional theory using expressions based on the harmonic transition state theory (for details see Refs. 18 and 23). By these expressions, the rate constants depend on the temperature T as well as on the CO and O₂ partial pressures p_{CO} and p_{O_2} , respectively. We will consider the reaction conditions of $T = 600$ K, $p_{\text{O}_2} = 1$ bar, and varying $p_{\text{CO}} \in [0.05, 50]$ bars, which result in the rate constants given in Table I. These are the reaction conditions, which we have employed in our previous studies^{3,16} and which therefore allow for direct comparison.

The TOF for the CO oxidation in the considered range of reaction conditions is shown in Fig. 1, i.e., the sum of the rates for the four CO₂ formation reactions normalized to the size of a surface unit cell. The data were obtained employing a lattice with 20×20 surface unit cells and by averaging over 10^9 kMC steps, after an initial relaxation of 10^8 steps. These are save settings, as we observe that stationary operation is usually reached after 10^6 steps. We find three regimes, which are characterized by their dominant site occupations: (i) at low p_{CO} , the surface is almost fully oxygen covered and the reactivity is very low (white background), (ii) the TOF is

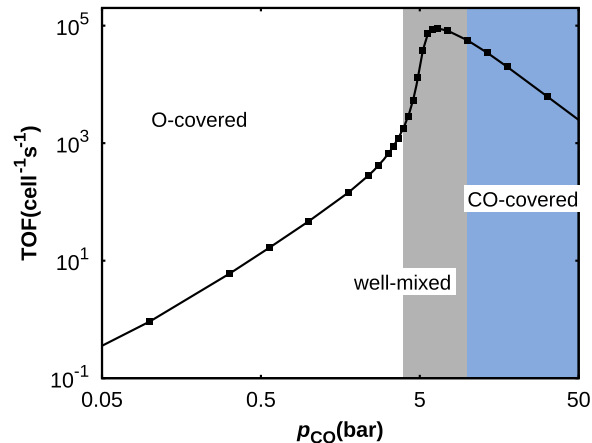


FIG. 1. Turnover frequency for the CO oxidation on RuO₂(110) for $T = 600$ K, $p_{\text{O}_2} = 1$ bar, and varying p_{CO} . At low p_{CO} , the surface is oxygen covered and the TOF is low (white background). With increasing partial pressure p_{CO} , the CO coverage increases as well. The TOF increases having its maximum, when both CO and oxygen are present in appreciable amounts (grey background). At high p_{CO} , the surface becomes CO-covered and the TOF decreases again (blue background).

the highest for medium p_{CO} , where both CO and O occupy the surface sites in roughly the same amounts (grey background), and (iii) the TOF decreases for higher p_{CO} as now the surface becomes CO poisoned (blue background). The reaction rates for all elementary steps can be found in the [supplementary material](#).

TABLE I. Elementary reaction steps and corresponding rate constants at the considered reaction conditions of $T = 600$ K, $p_{\text{O}_2} = 1$ bar, and varying $p_{\text{CO}} \in [0.05, 50]$ bars. An index *cus* or *br* indicates that the molecule is adsorbed at a cus- or br-site, respectively.

Process	ΔE^{\ddagger} (eV)	Rate constant (s ⁻¹)
Adsorption		
CO \rightarrow CO _{cus}	0.0	$2 \times 10^8 \times p_{\text{CO}}(\text{bars})$
CO \rightarrow CO _{br}	0.0	$2 \times 10^8 \times p_{\text{CO}}(\text{bars})$
O ₂ \rightarrow O _{cus} + O _{cus}	0.0	9.7×10^7
O ₂ \rightarrow O _{br} + O _{br}	0.0	9.7×10^7
O ₂ \rightarrow O _{br} + O _{cus}	0.0	9.7×10^7
Desorption		
CO _{cus} \rightarrow CO	1.3	9.2×10^6
CO _{br} \rightarrow CO	1.6	2.8×10^4
O _{cus} + O _{cus} \rightarrow O ₂	2.0	2.8×10^1
O _{br} + O _{br} \rightarrow O ₂	4.6	4.1×10^{-21}
O _{br} + O _{cus} \rightarrow O ₂	3.3	3.4×10^{-10}
Diffusion		
CO _{cus} \rightarrow CO _{cus}	1.7	6.6×10^{-2}
CO _{br} \rightarrow CO _{br}	0.6	1.1×10^8
CO _{cus} \rightarrow CO _{br}	1.3	1.5×10^2
CO _{br} \rightarrow CO _{cus}	1.6	0.5
O _{cus} \rightarrow O _{cus}	1.6	0.5
O _{br} \rightarrow O _{br}	0.7	1.6×10^7
O _{cus} \rightarrow O _{br}	1.0	4.9×10^4
O _{br} \rightarrow O _{cus}	2.3	6.0×10^{-7}
CO ₂ formation		
CO _{cus} + O _{cus} \rightarrow CO ₂	0.9	1.7×10^5
CO _{br} + O _{br} \rightarrow CO ₂	1.5	1.6
CO _{cus} + O _{br} \rightarrow CO ₂	1.2	5.2×10^2
O _{cus} + CO _{br} \rightarrow CO ₂	0.8	1.2×10^6

III. METHODS

We now turn to the problem of estimating the DRS from kinetic Monte Carlo simulations. As written above, we propose a three-step strategy for that purpose.

In the first step, we estimate bounds for the DRS using the relative entropy method as presented in Section III A. The ingredients to these bounds are the pathwise Fisher information metric and the time-integrated auto-correlation function (TAC) of the reaction rate. Especially, the latter is not trivial to estimate and, in Section III A 1, we present a sampling strategy based on a series expansion of a generalized inverse.

In the second step, we try to directly sample all DRSs from a single trajectory. For this, we employ our linear response theory based estimator, which we present in Section III B. The bounds from the first step help here to reduce the sampling overhead because we can skip those DRSs whose bounds are close to zero. We further utilize a numerical parameter, which we need to adjust during the determination of the TAC.

The direct estimator might suffer from a too high variance for some sensitivities and we would need unrealistically long trajectories to achieve a satisfactory accuracy. In the third step, we therefore determine these sensitivities using Coupled Finite Differences (CFDs),⁶ which might provide the desired variance with less sampling effort. However, we will need an extra kMC simulation for each DRS and if we can sufficiently determine at least some DRS from direct sampling, we will save computational costs. We outline how to implement CFDs for lattice problems in Section III C.

A. Bounds for the degree of rate sensitivity from the relative entropy method

As written above, this step's purpose is to reduce the computational costs at the later stages of the scheme. For this, we estimate upper bounds for all X_1^α and we can save CPU time by discarding those X_1^α which are close to zero. From this estimation process, we also obtain reasonable truncation limits for the next step in Section III B, the direct sampling of X_1^α based on the time-integrated linear response function.

For the bound, we will follow an approach put forward by Arampatzis, Katsoulakis, and Pantazis.^{17,22} For the DRS, their bound for a general parametric dependence specializes to

$$|X_1^\alpha| \leq \frac{1}{\langle R \rangle} \sqrt{c_R I^{\alpha\alpha}} =: B^\alpha, \quad (16)$$

where $I^{\alpha\alpha}$ are the diagonal elements of the pathwise Fisher Information Metric (FIM), and c_R is the time-integrated auto-correlation of the reaction rate. When the transition matrices $w_{ij}^\beta(\theta)$ arbitrarily depend on a parameter vector θ , the Fisher information metric is given by

$$I = \sum_{j \rightarrow i} P_{j,\text{stat.}} w_{ij} \nabla_\theta \log w_{ij} \nabla_\theta \log w_{ij}, \quad (17)$$

where the sum runs over all allowed events $j \rightarrow i$ ($w_{ij} \neq 0$). Using Eq. (3) and the logarithms of the rate constants as parameters, the above equation reduces to the reaction rates

$$I^{\alpha\beta} = \delta_{\alpha\beta} \langle R^\alpha \rangle. \quad (18)$$

So, using bound (16) will simply sort out those reactions which do not happen frequently enough. In the stationary case, the time-integrated auto-correlation (TAC) is

$$\begin{aligned} c_R &= 2 \int_0^\infty \langle \delta R(t) \delta R \rangle dt \\ &= \sum_{ij} 2 \int_0^\infty \delta R_i(e^{\Gamma t})_{ij} \delta R_j P_{j,\text{stat.}} dt, \end{aligned} \quad (19)$$

where $\delta R_i = R_i - \langle R \rangle$.

The bound (16) can be regarded as conservative sensitivity measures, i.e., only using the bounds we will definitely not miss an important elementary step. Even only considering the Fisher Information Metric (FIM) allows for a sensitivity analysis,²⁷ albeit with a different objective. The FIM measures the impact of a small change in the parameters on the whole stochastic process, i.e., the leading order of the relative entropy between the original and the perturbed processes. This allows for a global picture, but if one is interested in a particular average, the FIM based findings might be misleading. For example, we consider the case with two types of sites a and b , which are completely decoupled. Further, we want to assume that the reaction only takes place on sites of type a , but the processes on the site b are much faster. By Eq. (18), the processes on b -sites will have the largest FIM $I^{\alpha\alpha}$ and therefore the largest B^α . However, from the construction of our example, we know that these processes can have no impact on the average reaction rate.

1. An estimator for the time-integrated auto-correlation based on generalized inverses

With the definitions $v_i = \delta R_i$ and $u_i = \delta R_i P_{i,\text{stat.}}$, the TAC can be written as

$$\begin{aligned} c_R &= 2(v, \int_0^\infty e^{\Gamma t} dt u) = -2(v, \Gamma^\# u) \\ &= 2(v, D \sum_{l=0}^\infty (1 - \Gamma D)^l u), \end{aligned} \quad (20)$$

where (\cdot, \cdot) is the usual standard scalar product and $\Gamma^\#$ is a generalized inverse. The first line results from u and v being perpendicular to the left ($g_0 = (1, 1, \dots, 1)^T$), respectively, the right eigenvector ($g_i^0 = P_{i,\text{stat.}}$) to the eigenvalue zero. For the second line, we have employed the Neumann series for $\Gamma^\#$, where D is a suitable initial guess for $\Gamma^\#$. We choose

$$D_{ij} = \frac{1}{w_{ij}^{\text{acc.}}} \delta_{ij} = \Delta t_i \delta_{ij}, \quad (21)$$

which ensures the (linear) convergence of the series for the considered class of Markov processes (see the [supplementary material](#) for details). We can therefore truncate the series at some finite M (the truncation limit) and arrive at

$$\begin{aligned} c_R &\approx 2(v, D \sum_{n=0}^M (1 - \Gamma D)^n u) \\ &= 2 \sum_{ij} \delta R_i \Delta t_i \left(\sum_{n=0}^M (P_{\text{kMC}})^n \right)_{ij} \delta R_j P_{j,\text{stat.}}, \end{aligned} \quad (22)$$

where we have used $(1 - \Gamma D)_{ij} = P_{\text{kMC}}(i|j)$. Generally, a good choice of M is unknown and needs to be determined by testing the convergence.

As $(P_{\text{kMC}})^n_{ij}$ is the probability to be in the state i after n kMC steps starting in the state j , the TAC can be written as an expected value of a discrete time Markov chain, with the initial distribution $P_{i_0,\text{stat.}}$,

$$\begin{aligned} c_R &\approx 2 \sum_{i_0, i_1 \dots i_M} \left[\left(\sum_{n=0}^M \delta R_{i_n} \Delta t_{i_n} \delta R_{i_0} \right) \right. \\ &\quad \left. \times P_{\text{kMC}}(i_M | i_{M-1}) \dots P_{\text{kMC}}(i_1 | i_0) P_{i_0,\text{stat.}} \right]. \end{aligned} \quad (23)$$

We could thus run a kMC trajectory to obtain samples i_0 with their corresponding weights ($\propto \Delta t_{i_0}$) and simulating a single discrete time Markov chain for each initial i_0 . Fortunately, the latter is not necessary as the next M states in a kMC trajectory are generated according to P_{kMC} . We can save the time and just employ these. Putting this together with Eq. (15), we arrive at

$$c_R \approx \frac{2}{T} \sum_{l=M}^N \delta R_{i_l} \Delta t_l \sum_{n=0}^M \delta R_{i_{l-n}} \Delta t_{l-n}, \quad (24)$$

where $T = \sum_{l=0}^{N-M} \Delta t_l$.

An equation as (24) could also be motivated by first representing $\delta R(t)$ by a suitable trajectory averaging and then performing the (truncated) time integration. Choosing M large enough and a subsequent time step averaging results in the same formulas as Eq. (23). This allows us to interpret M as a measure

of the autocorrelation time in a number of kMC steps. Multiplying it with the average time step then leads to a physical autocorrelation time.

When implementing Eq. (24), the need to sum over the last M steps could become the computational burden, when M becomes large. A straightforward idea to lift this would be to employ a circular array which always stores $\delta R \Delta t$ of the last M steps and to introduce a variable S for the sum, which is updated in every step l according to the rule $S \rightarrow S - \delta R_{i_{l-M}} \Delta t_{l-M} + \delta R_{i_l} \Delta t_l$. Then the element in the circular array containing $\delta R_{i_{l-M}} \Delta t_{l-M}$ is overwritten by $\delta R_{i_l} \Delta t_l$. This allows an update in $\mathcal{O}(1)$ CPU-time and has $\mathcal{O}(M)$ storage requirement. However, for stiff problems as the CO oxidation on RuO₂(110), the time steps Δt_l easily vary by some orders of magnitude (and M can become large). We then end up in adding and subtracting small numbers from the variable S , which might quickly cause inaccuracies due to the finite machine precision. We therefore complement the circular array with a binary tree. The leaves then carry the elements of the circular array and an interior node stores the sum of the values stored at its children. The root node then carries the desired sum over all leaves. Choosing M as a power of two allows to employ a so-called perfect binary tree, i.e., every interior node has exactly two children and all leaves have the same depth.²⁸ This data structure and the corresponding operations can efficiently be implemented in a single array, which is of more relevance in kMC simulations than maximum flexibility in choosing M . If we now update the leaf carrying $\delta R_{i_{l-M}} \Delta t_{l-M}$, only the ancestor nodes need to be updated which scales $\mathcal{O}(\log_2 M)$. Being a little less efficient than the straightforward circular array, we, however, have the advantage, that at no point we need to subtract two numbers.

The [supplementary material](#) provides pseudocode for the sampling of the TAC during the kMC simulation.

B. Linear response theory based direct sampling of the DRS

The second step is the direct sampling of the sensitivities using our estimator based on the linear response theory. The linear response theory deals with small, in general time-dependent perturbations of the generator Γ in the master equation.²⁹ For our purposes, the perturbations result from changing the rate constants, i.e., we multiply the RC k^α with a factor $(1 + \varepsilon^\alpha(t))$. The perturbed generator is then given by

$$\Gamma_R(t) = \Gamma + \Gamma^\alpha \varepsilon^\alpha(t). \quad (25)$$

As we are interested in the stationary processes, we can restrict to the case, when the system initially was in the steady state for $\varepsilon^\alpha(t) = 0$, i.e., $P_j(t=0) = P_{j,\text{stat}}$. In the linear response regime, when $\varepsilon^\alpha(t)$ is small, the time dependent deviation $\delta \langle R \rangle(t)$ from the stationary expectation obeys²⁹

$$\langle R \rangle(t) - \langle R \rangle_{\text{stat}} = \int_0^t \chi^\alpha(t-s) \varepsilon^\alpha(s) ds, \quad (26)$$

where we omitted an explicit dependence of R on $\varepsilon^\alpha(t)$ as we target at an estimate for X_1^α . The Linear Response Function (LRF) $\chi^\alpha(t)$ in Eq. (26) is given by

$$\begin{aligned} \chi^\alpha(t) &= \sum_{ij} \delta R_i(t) \Gamma_{ij}^\alpha P_{j,\text{stat}} \\ &= \sum_{ijk} \delta R_i(e^{\Gamma t})_{ij} \Gamma_{ik}^\alpha P_{k,\text{stat}}, \end{aligned} \quad (27)$$

and can be obtained from the properties of the stationary initial state and the unperturbed process. If $\varepsilon^\alpha(t) = \text{const.}$ for $t > 0$, we expect that $\langle R \rangle(t)$ converges to a new steady state and, as we are in the linear regime with respect to the perturbation ε^α , we have

$$X_1^\alpha = \frac{1}{\langle R \rangle} \int_0^\infty \chi^\alpha(s) ds = (\tilde{R}, \Gamma^\# \Gamma^\alpha P_{\text{stat}}), \quad (28)$$

where $\tilde{R}_i = \delta R_i / \langle R \rangle$ and $\Gamma^\#$ is the same generalized inverse as in Eq. (20) because we have the same orthogonality properties for $\Gamma^\alpha P_{\text{stat}}$ as for $\delta R_i P_{i,\text{stat}}$. (compare Section III A 1 and the [supplementary material](#)).

1. Estimator for X_1^α

As in Section III A 1, we expand $\Gamma^\#$ into a series and arrive at

$$X_1^\alpha \approx \bar{X}_1^\alpha = \left(\tilde{R}, D \sum_{n=0}^M (P_{\text{kMC}})^n \Gamma^\alpha P_{\text{stat}} \right), \quad (29)$$

with D as defined by Eq. (21). We employ the same series expansion in the above equation as for the estimation of the TAC in Section III A 1. The truncation limit M should therefore be the same for both and we can employ a reasonable choice obtained in the first step also here.

Next, we decompose Γ^α into its diagonal part $\Gamma_{ij}^{\alpha,D}$ = $-\delta_{ij} \sum_l w_{ij}^\alpha$ and the off-diagonal partial transition matrix and make the decomposition $\bar{X}_1^\alpha = X_D^\alpha + X_W^\alpha$, where

$$\begin{aligned} X_W^\alpha \approx & \sum_{i_1, i_0, \dots, i_M} \left[\left(\sum_{n=0}^M \tilde{R}_{i_n} \Delta t_{i_n} O_{i_0, i_1}^\alpha \right) \right. \\ & \left. \times P_{\text{kMC}}(i_M | i_{M-1}) \dots P_{\text{kMC}}(i_0 | i_{-1}) P_{i_{-1}, \text{stat}} \right]. \end{aligned} \quad (30)$$

Here, we have used that $P_{\text{kMC}}(i|j) = 0 \Rightarrow w_{ij}^\alpha = 0$ and introduced the rescaled partial transition matrix O^α defined by

$$w_{ij}^\alpha = P_{\text{kMC}}(i|j) O_{ij}^\alpha, \text{ i.e., } O_{ij}^\alpha = \begin{cases} w_{ij}^{\text{acc.}}, & \text{if } j \rightarrow i \in \alpha \\ 0, & \text{else} \end{cases}. \quad (31)$$

Note, that in Eq. (30) X_W^α is an expected value over a discrete time Markov chain of length $M + 2$ instead of $M + 1$ as in Eq. (23). The contribution from the diagonal part of Γ^α can be written as

$$\begin{aligned} X_D^\alpha \approx & \sum_{i_1, i_0, \dots, i_M} \left[\left(\sum_{n=-1}^M \tilde{R}_{i_n} \Delta t_{i_n} w_{i_{-1}, i_1}^{\alpha, \text{acc.}} \right) \right. \\ & \left. \times P_{\text{kMC}}(i_M | i_{M-1}) \dots P_{\text{kMC}}(i_0 | i_{-1}) P_{i_{-1}, \text{stat}} \right], \end{aligned} \quad (32)$$

where we have added a term $(\tilde{R}, D(1 + \Gamma D)^{M+1} \Gamma^{\alpha,D} P_{\text{stat}})$, so that the Markov chains in Eqs. (30) and (32) have the same length. We can do so as this term converges to zero for large enough M (see the [supplementary material](#)).

Putting Equations (30) and (32) together and employing the same argumentation as in Section III A 1, we arrive at the estimator

$$X_1^\alpha \approx \frac{1}{T} \sum_{n=M+1}^N \left[-\tilde{R}_{i_n} w_{i_n}^{\alpha, \text{acc.}} \Delta t_n^2 + \tilde{R}_{i_n} \Delta t_n \sum_{l=1}^{M+1} (O_{i_{n+1}i_n}^\alpha - w_{i_{n-l}}^{\alpha, \text{acc.}}) \Delta t_{n-l} \right], \quad (33)$$

where $T = \sum_{l=0}^{N-M-1} \Delta t_l$. We term this way to estimate X_1^α Integrated Response Function (IRF) approach. Although being based on a very different derivation, the IRF estimator (33) is very similar to the estimator presented in Refs. 10 and 21, except that it employs deterministic Δt instead of random time steps δt . These deterministic Δt might help to improve the performance, as it is the case for log-likelihood estimators.^{11,12} However, the motivation from generalized inverses and series expansion might open ways to improve its performance, e.g., by choosing a different initial guess D .

The above formula (33) explains why we have added the extra term in Eq. (32): Both $O_{i_{n+1}i_n}^\alpha$ and $w_{i_n}^{\alpha, \text{acc.}}$ can now be calculated at the kMC step n . The accumulated rate needs to be calculated anyways and for $O_{i_{n+1}i_n}^\alpha$ we just need to know which type of reaction will be executed next. We can therefore employ the same binary tree for summation as for the TAC sampling with negligible overhead. The leaves will now carry the values $(O_{i_{n+1}i_n}^\alpha - w_{i_n}^{\alpha, \text{acc.}}) \Delta t_n$ instead of $\delta R_{i_n} \Delta t_n$.

Pseudocode for the sampling of the IRF can be found in the [supplementary material](#).

C. Coupled finite differences

The direct sampling of the sensitivities is not always successful. For some DRSS, we might still have too large sampling errors. We estimate these DRSS, which are not sufficiently accurate, using Finite Difference Approximations (FDAs). In particular, we employ central finite differences,

$$\left(\frac{\partial \langle R \rangle (\{k^\beta\})}{\partial \log k^\alpha} \right)_{k^\beta \neq \alpha} \approx \frac{\langle R \rangle ((1+h)k^\alpha) - \langle R \rangle ((1-h)k^\alpha)}{2h}, \quad (34)$$

which are second order accurate in the difference parameter h . Straightforward FDA uses two independent kMC simulations to estimate $\langle R \rangle (\dots (1+h)k^\alpha \dots)$ and $\langle R \rangle (\dots (1-h)k^\alpha \dots)$. As h approaches zero, the difference between $\langle R \rangle (\dots (1+h)k^\alpha \dots)$ and $\langle R \rangle (\dots (1-h)k^\alpha \dots)$ will become smaller than the sampling errors in the statistically independent estimates, and the FDA might thereby carry a huge error. To overcome this to a certain degree, the two simulations could be coupled such that the fluctuations of the two estimates point in the same direction and the estimated difference has a smaller sampling error than the two estimates for $\langle R \rangle$. The simplest of such couplings is the common random number approach,³⁰ where we would employ the same set of pseudo random numbers for both kMC simulations. This approach is easy to implement, but seems not to provide significant improvements for lattice systems.⁷ More advanced coupling strategies exist (see, e.g., Ref. 7), of which

we choose the Coupled Finite Differences (CFDs) proposed by Anderson,⁶ for the ease of its implementation.

To outline how CFDs can be applied to lattice problems, it is convenient to decompose the set β into different reaction channels. That is, two transitions $i_1 \rightarrow j_1$ and $i_2 \rightarrow j_2$ belong to the same reaction channel A_β , if they belong to the same reaction β and cause the same change $\xi^{A_\beta} := j_1 - i_1 = j_2 - i_2$. Intuitively, this introduces a spatial resolution of the reaction set β . While β includes, for instance, all CO adsorption on cus sites, a CO adsorption on the first cus site will not belong to the same reaction channel as a CO adsorption on the second cus site, simply because both change different entries in the integer vector representing the current state of the system. In contrast, all transitions, which correspond to a CO adsorption on the first cus site, will be in the same channel as they all change the same entry from *empty* to *CO* and leave all others unchanged. With this concept, we can rewrite the master equation as

$$\frac{d}{dt} P_i(t) = \sum_{\beta} \sum_{A_\beta} \left[a^{A_\beta}(i - \xi^{A_\beta}) P_{i - \xi^{A_\beta}}(t) - a^{A_\beta}(i) P_i(t) \right], \quad (35)$$

where, in the context outlined in Section II, the reaction propensity $a^{A_\beta}(i)$ has the form

$$a^{A_\beta}(i) = \begin{cases} k^\beta, & \text{if } i \rightarrow i + \xi^{A_\beta} \in A_\beta. \\ 0, & \text{else.} \end{cases} \quad (36)$$

The two independent trajectories for an uncoupled finite difference approximation can now be regarded as a single process operating on two lattices. The corresponding master equation is then

$$\begin{aligned} \frac{d}{dt} P_{i,j}(t) = & \sum_{\beta} \sum_{A_\beta} \left[a_h^{A_\beta}(i - \xi^{A_\beta}) P_{i - \xi^{A_\beta}, j}(t) - a_h^{A_\beta}(i) P_{i,j}(t) \right] \\ & + \sum_{\beta} \sum_{A_\beta} \left[a_{-h}^{A_\beta}(j - \xi^{A_\beta}) P_{i, j - \xi^{A_\beta}}(t) - a_{-h}^{A_\beta}(j) P_{i,j}(t) \right], \end{aligned} \quad (37)$$

where $P_{i,j}(t)$ is the joint probability to find the first lattice in the state i and the second lattice in the state j . The index h at the propensity $a_h^{A_\beta}(\cdot)$ indicates that this is obtained by using the RCs $\{\dots, (1+h)k^\alpha, \dots\}$. Here k^α is the RC, for which we want to estimate the corresponding DRS. Summing over j will therefore lead to the master equation (35) for the RCs $\{\dots, (1+h)k^\alpha, \dots\}$ and summing over i will achieve the same for $\{\dots, (1-h)k^\alpha, \dots\}$. Correspondingly, $\langle R \rangle ((1+h)k^\alpha)$ will be estimated by only using the i -component of the joint process and $\langle R \rangle ((1-h)k^\alpha)$ by only using the j -component. If we now calculate the variance of the difference $R_h(i) - R_{-h}(j)$ using the joint probability $P_{i,j}$, we obtain

$$\text{Var}(R_h - R_{-h}) = \text{Var}(R_h) + \text{Var}(R_{-h}) - 2\langle \delta R_h \delta R_{-h} \rangle, \quad (38)$$

i.e., the variance of the difference would be reduced if there would be a positive correlation $\langle \delta R_h \delta R_{-h} \rangle$ between $R_h(i)$ and $R_{-h}(j)$. For the uncoupled FDA obeying Eq. (37) the correlation $\langle \delta R_h \delta R_{-h} \rangle$ is, of course, zero due to the statistical independence of i and j . The idea is now to introduce a new joint process, which gives the same averages $\langle R_h \rangle$ and $\langle R_{-h} \rangle$ but a positive correlation $\langle \delta R_h \delta R_{-h} \rangle$. There are multiple different possibilities to realize this.⁷ Anderson's method employs a process obeying the master equation

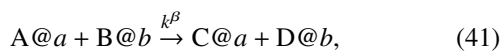
$$\begin{aligned}
\frac{d}{dt}P_{i,j}(t) &= \sum_{\beta} \sum_{A_{\beta}} \left[a_1^{A_{\beta}}(i - \xi^{A_{\beta}}, j - \xi^{A_{\beta}}) P_{i - \xi^{A_{\beta}}, j - \xi^{A_{\beta}}}(t) - a_1^{A_{\beta}}(i, j) P_{i,j}(t) \right] \\
&+ \sum_{\beta} \sum_{A_{\beta}} \left[a_2^{A_{\beta}}(i - \xi^{A_{\beta}}, j) P_{i - \xi^{A_{\beta}}, j}(t) - a_2^{A_{\beta}}(i, j) P_{i,j}(t) \right] \\
&+ \sum_{\beta} \sum_{A_{\beta}} \left[a_3^{A_{\beta}}(i, j - \xi^{A_{\beta}}) P_{i, j - \xi^{A_{\beta}}}(t) - a_3^{A_{\beta}}(i, j) P_{i,j}(t) \right]
\end{aligned} \quad (39)$$

with the propensities,

$$\begin{aligned}
a_1^{A_{\beta}}(i, j) &= \min(a_h^{A_{\beta}}(i), a_{-h}^{A_{\beta}}(j)) \\
&= \begin{cases} k_{-h}^{\beta} & \text{if } i \rightarrow i + \xi^{A_{\beta}} \in A_{\beta}, \\ & \text{and } j \rightarrow j + \xi^{A_{\beta}} \in A_{\beta}, \\ 0 & \text{else,} \end{cases} \\
a_2^{A_{\beta}}(i, j) &= a_h^{A_{\beta}}(i) - \min(a_h^{A_{\beta}}(i), a_{-h}^{A_{\beta}}(j)) \\
&= \begin{cases} k_h^{\beta} & \text{if } i \rightarrow i + \xi^{A_{\beta}} \in A_{\beta}, \\ & \text{and } j \rightarrow j + \xi^{A_{\beta}} \notin A_{\beta}, \\ k_h^{\beta} - k_{-h}^{\beta} & \text{if } i \rightarrow i + \xi^{A_{\beta}} \in A_{\beta}, \\ & \text{and } j \rightarrow j + \xi^{A_{\beta}} \in A_{\beta}, \\ 0 & \text{else,} \end{cases} \quad (40) \\
a_3^{A_{\beta}}(i, j) &= a_{-h}^{A_{\beta}}(j) - \min(a_h^{A_{\beta}}(i), a_{-h}^{A_{\beta}}(j)) \\
&= \begin{cases} k_{-h}^{\beta} & \text{if } i \rightarrow i + \xi^{A_{\beta}} \notin A_{\beta}, \\ & \text{and } j \rightarrow j + \xi^{A_{\beta}} \in A_{\beta}, \\ 0 & \text{else,} \end{cases}
\end{aligned}$$

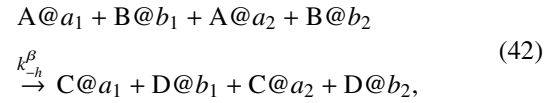
where we inserted Eq. (36) for the respective second equal sign. Using that, we can split the propensity $a_2^{A_{\beta}}(i, j)$ into two new ones $a_{2,1}^{A_{\beta}}(i, j)$ and $a_{2,2}^{A_{\beta}}(i, j)$. The first will be non-zero ($a_{2,1}^{A_{\beta}}(i, j) = k_h^{\beta}$) only when the first case fulfilled ($i \rightarrow i + \xi^{A_{\beta}} \in A_{\beta}$ and $j \rightarrow j + \xi^{A_{\beta}} \notin A_{\beta}$). The second will only be non-zero ($a_{2,2}^{A_{\beta}}(i, j) = k_h^{\beta} - k_{-h}^{\beta}$) for the second case ($i \rightarrow i + \xi^{A_{\beta}} \in A_{\beta}$ and $j \rightarrow j + \xi^{A_{\beta}} \in A_{\beta}$).

We can describe the Markov process defined by Eqs. (39) and (40) in a more chemical language, which allows an implementation using the standard interfaces of lattice kMC codes. If we have the sites (a, b, c, \dots) for our original problem, the process for CFD operates on a lattice with the sites ($a_1, b_1, c_1, \dots, a_2, b_2, c_2, \dots$). For the CO oxidation on RuO₂(110), this can be achieved by replacing the two site types (br and cus) by the four sites (br₁, cus₁, br₂, and cus₂). Suppose the reaction channel A_{β} originally corresponds to a change of the adsorption states on the sites (a, b), i.e., it can be represented by

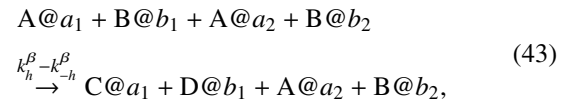


where $A@a$ means that the site a carries the adsorbate A , which is altered to the adsorbate C after the process has been executed. On our doubled lattice for CFD, we have to replace this with appropriate new reaction channels. The reaction channel corresponding to $a_1^{A_{\beta}}$ simultaneously executes the original process

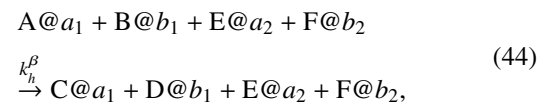
(41) on both lattices, i.e., we have to introduce the reaction



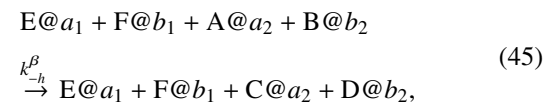
where $\xrightarrow{k_{-h}^{\beta}}$ indicates that this will be executed with a rate constant k_{-h}^{β} . Correspondingly, the reaction channel for $a_{2,2}^{A_{\beta}}$ will have the reaction equation



as it only changes the configuration on the first lattice, but requires this change to be possible on the second lattice. What remains are the channels, which correspond to $a_{2,1}^{A_{\beta}}$ and $a_3^{A_{\beta}}$. These change the configuration on one lattice according to A_{β} and require that A_{β} is not possible on the respective other lattice. The required logic is usually not implemented in standard interfaces of lattice kMC codes (at least, kmcs does not allow this). We therefore employ the same trick which leads to the splitting of $a_2^{A_{\beta}}$ and introduce a new reaction channel for each situation which fulfills the requirements. In other words, we introduce for every pair (E, F) $\neq (A, B)$ the reaction



which corresponds to $a_{2,1}^{A_{\beta}}$ and the reaction



which corresponds to $a_3^{A_{\beta}}$. We can thus employ the standard interface of kmcs for defining the process and perform the CFD without needing to touch the source code. The price we pay is that for a two-site reaction as (41) in the original process, we have to introduce $\mathcal{O}(N_{\text{spec}})^2$ in the process for the CFD. Here, N_{spec} is the number of surface species or more general the number of states a site can attain. Probably, this is not the most efficient way of implementing CFD and more complex reactions operating on more than two sites will require even more additional reactions. However, for the case at hand, this overhead stays reasonable and running N kMC steps for the CFD comes at roughly five times the cost of running N kMC steps of the original process.

IV. RESULTS

As described in Section II D, we study the CO oxidation on the RuO₂(110) surface for the reaction conditions $T = 600$ K, $p_{\text{O}_2} = 1$ bar, and $p_{\text{CO}} \in [0.05, 50]$ bars. The investigated reaction rate is the CO oxidation turnover frequency (displayed in Figure 1) and we test its sensitivity to all (22) different rate constants using the outlined three-stage strategy.

A. First step: Bounds

The first step is the estimation of bounds for the DRS from the Fisher Information Metric (FIM) and the time-integrated autocorrelation function (TAC). In all simulations, we employ lattices of 20×20 unit cells as in previous studies.^{16,18} After an initial relaxation to the steady state with 10^8 steps, the TAC is sampled from 100 sub-trajectories each having a length of 10^7 steps and using 10^7 decorrelation steps between the sampling trajectories.

In practice, a good choice for M is usually unknown and thus we have to test different M for estimating the TAC c_R until we reach convergence. As the choice of M does not alter the kMC code, we can perform this test for a range of possible values during a single kMC simulation. Although we do not expect small values of M to be sufficient, we test all powers of two between $2^0 = 1$ and $2^{20} \approx 10^6$. For the case $p_{\text{CO}} = 1$ bar, we show the dependence of the TAC estimates on the choice of M in Figure 2. Also shown are the standard deviations of the TAC estimates as error bars. For small values of M , these are very small but increase for larger values of M . Thus the choice of M has to balance two errors. It needs to be large enough such that expansion (22) is accurate, and, on the other hand, small enough such that the variance of the sampled estimate does not grow too large. This can be seen in Figure 2, where the TAC is well converged for a choice $M = 2^{12} = 4096$, and the higher expansion accuracy for larger M gets corrupted by the sampling error.

All DRSs must add up to one and we therefore expect the relevant DRSs to be in this order of magnitude (see the [supplementary material](#) for the proof). As a threshold for the relevance, we choose $tol = 0.02$, i.e., all DRSs with a bound B^α smaller than

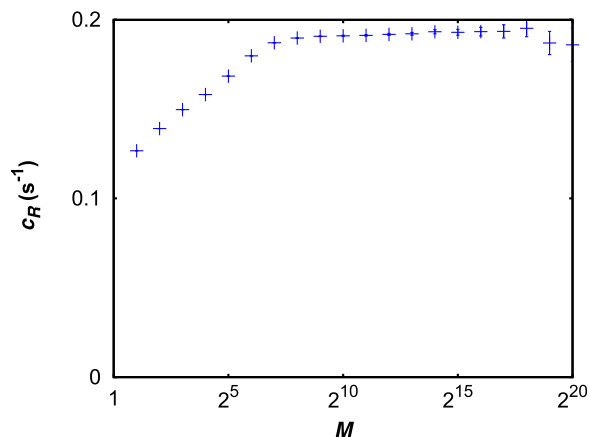


FIG. 2. Estimate of the time-integrated auto-correlation function (TAC) as a function of the expansion parameter M for the CO oxidation on RuO₂(110) for $T = 600$ K, $p_{\text{O}_2} = 1$ bar, and $p_{\text{CO}} = 1$ bar. Also shown is the standard deviation as error bars. The TAC is well converged for a choice $M = 2^{12} = 4096$.

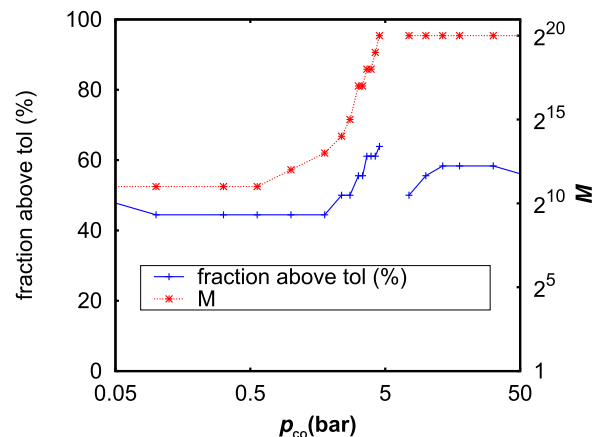


FIG. 3. Fraction of all reactions with a sensitivity bound B^α above 2% (blue, left axis) and the corresponding value for the expansion parameter M (red, right axis).

this will not be calculated later on. Figure 3 shows the fraction of all DRSs, whose bounds B^α are above the threshold tol (blue, left axis) in the dependence of p_{CO} and for fixed $p_{\text{O}_2} = 1$ bar and $T = 600$ K. Further, we indicate our choice for the truncation limit M (red, right axis). Notably, we found no convergence for $M \leq 2^{20}$ for CO partial pressures around 5 bars, i.e., in the vicinity of the point of highest reactivity. The interpretation of the expansion (22) as a time integration suggests that we are close to a second order phase transition, where the correlation time diverges and thus the integration limit must be chosen very large. A closer look at the coverage curves in Ref. 2 corroborate this interpretation, but also the results on spatial correlation from Refs. 19 and 20 point in this direction. Also, we can work with much smaller truncation limits M for small p_{CO} than for large, although, in both cases, we are not close to the reaction conditions for which we expect a second order phase transition. We explain this as follows. For low p_{CO} , in the O-covered regime, the kinetics almost exclusively takes place on the cus sites¹⁶ and the surface lattice is fully covered with oxygen. An adsorption of CO onto a vacant site, resulting from an oxygen desorption, triggers with very high probability either a CO desorption or a CO+O reaction in the next step, as the RCs for both are much higher than for any other possible reactions. The resulting vacancies are likely to be filled with oxygen (due to the high O₂ adsorption RC). We thus quickly return to a fully oxygen covered surface. We therefore expect the correlation time in a number of kMC steps to be rather short. For high p_{CO} and the concomitant almost always fully CO-covered surface, the removal of oxygen would require much longer as now there are many fast competing processes. For instance, CO desorption from cus sites has a ten times higher RC than the most likely O_{cus} + CO_{br} reaction and there are many CO_{cus}. Thus the decay of the oxygen adsorbed state would take rather many kMC steps.

In all cases, we found that roughly 50% of the DRS bounds are above our tolerance. For the case, that we would directly proceed with a finite difference approximation as in Ref. 22, this would translate into a 50% saving in CPU-time. In our approach, we add the direct estimation of the sensitivities between the bound estimation and the Finite Difference Approximation. The question naturally arises, what is the benefit of the bound estimation with an extra kMC simulation to

save 50% of the sampling effort in the subsequent kMC simulation for the direct estimation of the sensitivities. The answer is that this step is not only there for ruling out DRSs but also to identify a reasonable choice of M as, in Sec. IV B, we will employ the same expansion which led to Eq. (22). Thus, this pre-screening step reduces the sampling effort to $50\%/N_M$ as we save the examination of N_M different values of M . As we need to sample only a single TAC but N_{reac} DRSs, determining M in this step allows for significant savings in the subsequent step.

B. Second step: Direct sampling of the DRS

As in Section IV A, we consider first the case $T = 600$ K, $p_{\text{O}_2} = 1$ bar, $p_{\text{CO}} = 1$ bar, and the dependence on the expansion length M . For this, we ignore the results from the bound estimations and sample all DRSs. In order to test the computational costs, we, however, perform an extra kMC simulation for each M . For these, we employ similar settings as in Section IV A, i.e., a 20×20 lattice and sampling over 100 snippets of 10^7 kMC steps, between which we perform 10^6 steps for decorrelation. As in Section IV A, we employ these 100 samples to obtain the estimated DRSs and corresponding variances. The results are summarized in Figure 4. As this is the demanding task, we only show X_1^α in the top panel. The estimates are well converged for $M = 2^{12}$, i.e., the value we have extracted from the simulation of the TAC. The variance of the estimation linearly increases with M , which is reflected by the standard deviation, shown in the panel below. For the current set of reaction conditions and a reasonable choice of M , the standard deviations are in the order of 10^{-2} or below, so that we estimate all DRSs

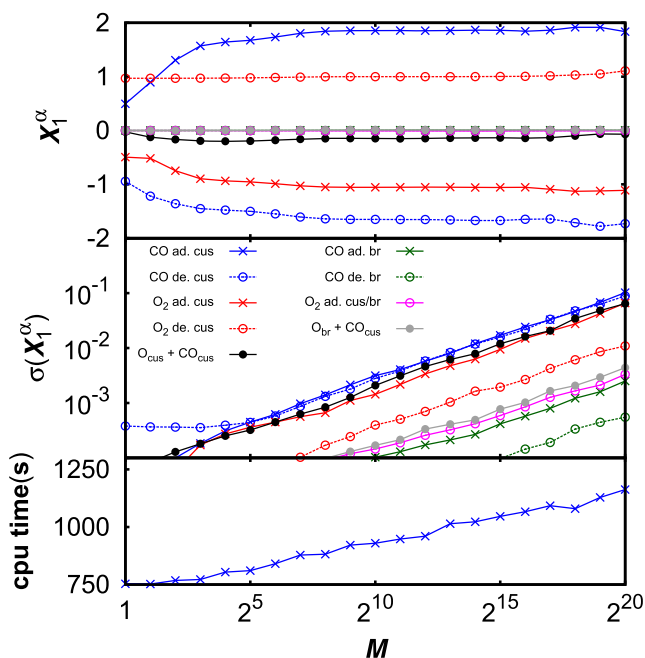


FIG. 4. Estimates of the Degree of Rate Sensitivities (DRSs) based on the linear response sampling as a function of the expansion parameter M for the CO oxidation on $\text{RuO}_2(110)$ for $T = 600$ K, $p_{\text{O}_2} = 1$ bar, and $p_{\text{CO}} = 1$ bar (top panel). As the TAC (see Fig. 2), the DRSs are converged for a choice $M = 2^{12} = 4096$. The standard deviations of the sampled DRSs grow with the root of M (middle panel). The lower panel shows the CPU time, which grows, as expected, linear with the logarithm of M (up to an additive constant).

for the cost of two extra kMC simulations (one for estimating the bounds and M and one for estimating the DRSs). The CPU-time for the last simulation comes at only slightly higher costs by the extra sampling as can be seen in the lowest panel in Fig. 4. As expected, the sampling overhead linearly depends on $\log M$ and, for $M = 2^{20} \approx 10^6$, it roughly requires one third of the CPU time. The computational savings of the two-step procedure stem from the possibility to avoid the test of different values of M . Without this prescreening of different M values, the sampling overhead would be responsible for $>80\%$ of the CPU load for the considered case. The simulation would therefore run roughly six times as long as a pure kMC simulation. Utilizing the estimated bounds obtained in Section IV A, the overhead could have been further reduced by roughly 50%. The sampling of the TAC comes at roughly the same costs as a full DRS sampling at fixed M , as we have to test for multiple M . So, we end up with an overhead of a factor 2-3 (including the CPU-time for the prescreening). Without prescreening we instead would arrive at an overhead of a factor ≈ 6 . For more complicated reaction mechanism and the concomitant larger number of RCs, this saving would increase.

With the estimated bounds and the values for M obtained in Section IV A, we now sample the DRSs for varying CO partial pressures. We estimate these from 1000 sub-trajectories each of the size $100 \times M$ steps, but at least 10^6 and at most 10^7 . As M can be regarded as a correlation time, we employ $10 \times M$ decorrelation steps between the samples. As for Fig. 3, we do not perform an estimation around the point of highest activity, since we would need M to be larger than one million. In principle, larger M could have been employed, but, as the

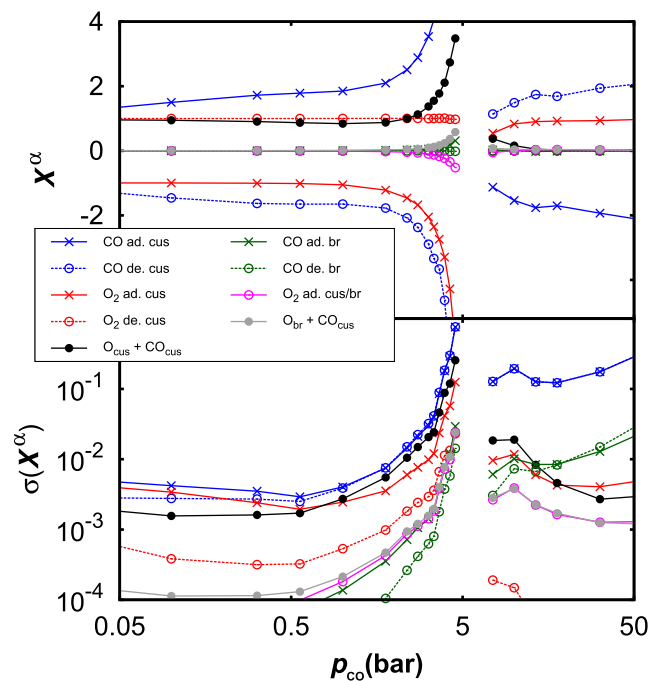


FIG. 5. The sampled Degree of Rate Sensitivity (DRS) based on the linear response sampling as a function of the CO partial pressure p_{CO} ($T = 600$ K, $p_{\text{O}_2} = 1$ bar, top panel). The fluctuations during sampling expressed as standard deviation σ are shown in the lower panel. The simulation employed the expansion parameters M obtained in Sec. IV A (compare Fig. 3). Not shown are those DRSs which are significantly below zero and the range around $p_{\text{CO}} = 5$ bars, for which $M > 2^{20}$ would be required.

variance linearly increases with M , we do not expect accurate enough estimates.

The results for the DRSs are shown in the top panel of Fig. 5. Shown are only those DRSs, which are non-zero in the range of the considered reaction conditions. In most cases, the number of non-vanishing DRSs is small, even considering the pre-screening. Inspecting the bottom panel of Fig. 5, we find that the standard deviation of our estimates largely depends on the considered reaction. Especially, those DRSs, which are low, can be estimated pretty well with very small sampling errors. However, also most other DRSs can be obtained with an accuracy below 0.1, i.e., reasonably accurate for many purposes. Only close to the gap and for the CO ad/desorption on cus sites at high CO partial pressures, the standard deviations exceed the value of 0.1. This increase in the standard deviations directly corresponds to the larger values of M required there. For high p_{CO} , we observe that the DRS estimates for the fastest processes (CO ad/desorption on cus sites) have the highest variance. This behaviour is similar to the Girsanov transform based approach.¹² If we need more accurate estimates with $\sigma(X_1^\alpha) < 0.1$, we could increase the sampling time. However, the sampling error behaves as $\mathcal{O}(N^{-\frac{1}{2}})$, where N is the total number of kMC steps. So pushing the error below the targeted limit might become very demanding for those DRSs, which cannot be sampled well, and it might become reasonable to employ a different strategy for obtaining these.

C. Third step: Coupled finite differences

We now employ CFD to estimate the DRSs, which could not be sampled with sufficient accuracy by the direct sampling approach from Sec. IV B. These are all DRSs which lay in the gap around $p_{\text{CO}} = 5$ bars and those for which the existing estimates have a standard deviation above 0.1. To estimate the DRSs and the respective variances, we employ 1280 samples each having 10^6 kMC steps. All other settings are as in Section IV B. For the difference parameter, we choose $h = 0.01$.

Figure 6 shows the final results combining the IRF and CFD estimates (top panel) with the corresponding standard deviation (lower panel). Symbols mark data points which have been obtained using the CFD, all others have directly been sampled by the IRF. The standard deviations for the estimates are now below the desired value of 0.1, except close to the point of highest activity where, for some DRSs, we stay a little above the desired accuracy. However, the affected DRSs are rather large there and the relative error is of an acceptable order.

Having a closer look, we find that for the most of the pressure range in Fig. 6 we just need a few or no CFD estimates. This especially holds true for low p_{CO} where all DRSs can accurately be estimated and no CFD needs to be performed. This can be explained by the relatively small values for the expansion order M in the IRF, as the variance of the estimator grows linearly with M and thus small M should rather lead to accurate estimates. But also for large p_{CO} only the DRSs for CO ad/desorption on cus sites needed improvements. Only close to the pressures where the DRSs for CO and O₂ ad/desorption on cus sites “diverge” multiple DRSs needed refinements using CFD.

The procedure, as outlined, leaves room for improvement if CPU time really is the limiting factor. Most notably, we

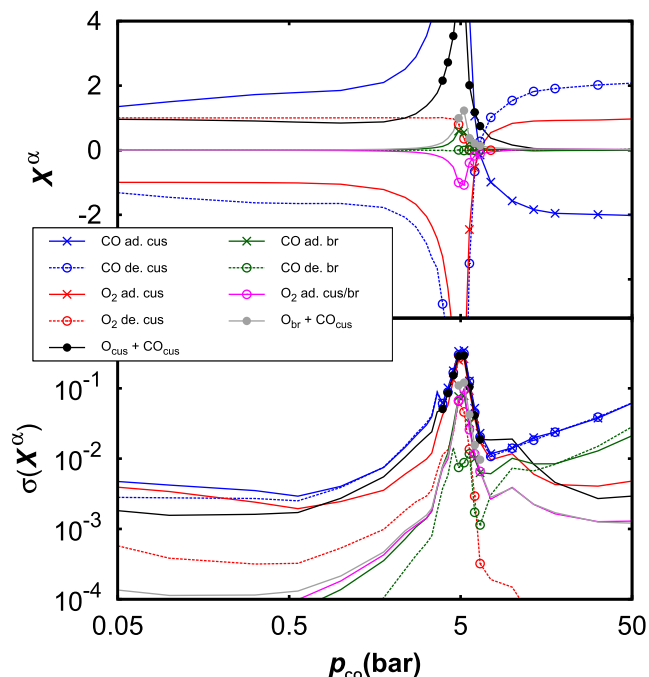


FIG. 6. Same as Fig. 5, but missing points (around $p_{\text{CO}} = 5$ bars) and those with a too high variance have been determined using coupled finite differences. These data are represented using symbols; all other points have been obtained from the linear response sampling.

could perform the IRF with much less kMC steps. At low p_{CO} , the DRSs have an accuracy which is roughly ten times lower than required, so we could save a factor around hundred here. Also, we would not need to perform ten billion kMC steps for the IRF sampling at high CO partial pressures. Except for CO ad/desorption on cus sites, all DRSs would have the desired accuracy already at around one billion steps. Also running IRF with $M > 10^6$ in the gap around $p_{\text{CO}} = 5$ bars might not help a lot for the large DRSs, but it might allow to determine the lower ones with sufficient accuracy. However, even in this unoptimized form, we significantly reduce the computational effort with respect to the very expensive numerical derivatives.

V. DISCUSSION

We now turn to the discussion of the results obtained in Sec. IV. As it is the more common measure, we calculated the Degree of Rate Control (DRC) from our DRS results and Eq. (14). The DRC is shown in Fig. 7. For low CO partial pressures, the DRC identifies only a single rate-determining step, which is the CO_{cus} + O_{cus} reaction. For high CO partial pressures, the oxygen adsorption on cus-sites is the only rate-determining step. Between these two extremes, we find multiple steps, which have significant DRCs. Close to 5 bars, the absolute DRCs of all displayed reactions are in the order of one or larger and the notion of a single rate-determining step seems not appropriate, anymore.

If we compare the DRC in Fig. 7 with the DRS in Fig. 6, the DRC provides a clearer picture. In contrast, the DRS plot is busier but also provides information which is not present in the DRCs. At low p_{CO} , the DRS identifies adsorption and desorption of CO and oxygen on cus-sites to have an impact on the TOF, with opposite signs but (roughly) the same absolute values

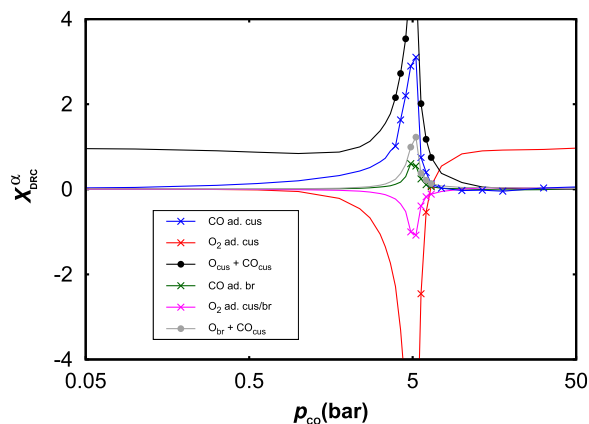


FIG. 7. The sampled degree of rate control as a function of the CO partial pressure p_{CO} ($T = 600$ K, $p_{\text{O}_2} = 1$ bar, top panel). Symbols represent data which were obtained using coupled finite differences.

for forward and backward reactions. This allows for an intuitive microscopic interpretation. At these conditions, the surface is fully oxygen covered and we need to get CO onto the surface for a high TOF with the employed Langmuir-Hinshelwood mechanism. To achieve this, empty sites need to be created by one of the oxygen desorption processes and the most effective $2\text{O}_{\text{cus}} \rightarrow \text{O}_2(\text{gas})$ desorption has a positive DRS. Now, CO and oxygen adsorptions onto the created empty sites compete with each other. So, the $\text{O}_2(\text{gas}) \rightarrow 2\text{O}_{\text{cus}}$ adsorption has a negative DRS as it blocks the more catalytically active configuration. The CO adsorption on cus sites has a positive sign because it leads to configurations for which the $\text{O}_{\text{cus}} + \text{CO}_{\text{cus}} \rightarrow \text{CO}_2(\text{gas})$ reaction is possible. Once a CO molecule is adsorbed on a cus site, we have a competition between the desorption and the actual reaction. The desorption destroys the desired configuration and therefore has a negative sign. The $\text{O}_{\text{cus}} + \text{CO}_{\text{cus}} \rightarrow \text{CO}_2(\text{gas})$ has a positive sign, which now results from the two competing parts X_0^α and X_1^α . The direct relative contribution X_0^α to the TOF is always positive. At these low CO partial pressures, the $\text{O}_{\text{cus}} + \text{CO}_{\text{cus}} \rightarrow \text{CO}_2(\text{gas})$ reaction is responsible for almost all the reactivity and its relative contribution is therefore very close to one. This can be seen from Fig. 8, which shows the sensitivity bounds B^α (Eq. (16)), which are proportional to the square roots of reaction rates of each elementary step. The second is the contribution X_1^α to the creation/annihilation of the desired configurations, which should be negative for this case because the reaction removes the CO from the surface. This contribution is close to zero as the dominant channel of the CO removal from the surface is the CO desorption from cus sites, which has a higher frequency (compare Fig. 8).

At low CO partial pressures, the cancellation of the DRS for ad/desorption on cus sites suggests that these steps are close to equilibrated. This interpretation is supported by the bounds for the forward and backward rates, which are almost the same for CO ad/desorption and oxygen ad/desorption on cus sites, respectively. As we increase p_{CO} , the CO coverage on cus sites also increases. Although this coverage is still very low, the fast $\text{O}_{\text{cus}} + \text{CO}_{\text{cus}} \rightarrow \text{CO}_2(\text{gas})$ reaction becomes an alternative to the O_{cus} -desorption for oxygen removal from the surface. This is reflected by the deviation of the bounds for O_{cus} -desorption and O_{cus} -adsorption in Fig. 8. In consequence, the DRC for the

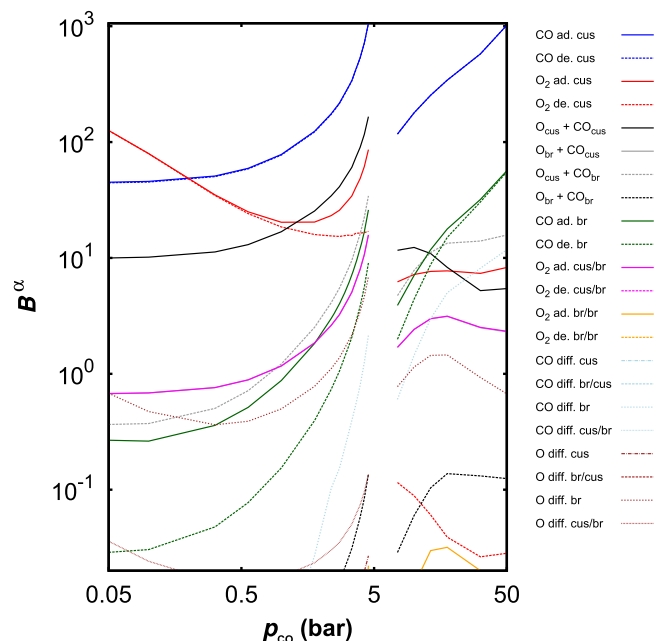


FIG. 8. Bounds B^α for the degree of rate sensitivity as a function of the CO partial pressure p_{CO} ($T = 600$ K, $p_{\text{O}_2} = 1$ bar, top panel). The bounds show the same rapid increase close to $p_{\text{CO}} = 5$ bars as the DRS. However, some elementary steps would be marked important by the bounds, but are not according to the DRS criterion, e.g., adsorption and desorption of CO on bridge sites.

pair of O_{cus} -ad/desorption starts deviating from zero, as they are not equilibrated anymore. For the point at $p_{\text{CO}} = 5$ bars, we speculated that its very close to a second order phase transition and the correlation time as well as fluctuations should increase as we approach this point. This results in an increase of the time-integrated auto-correlation (TAC), which is reflected in an increase of the bounds for most elementary steps. The same holds for most integrated response functions X_1^α , and, in consequence, the absolute DRS for the relevant elementary steps increases while approaching the “critical point.” For the DRC in Fig. 8 for the CO_{cus} -ad/desorption, this results in an amplification of the effect of the slight deviation from equilibrium. The O_{cus} -ad/desorption is much better equilibrated at low p_{CO} and therefore this amplification of the DRC sets in at little higher p_{CO} . In contrast to the CO_{cus} -ad/desorption, the increase in the absolute DRC of the O_{cus} -ad/desorption is more pronounced because this reaction pair is additionally strongly driven out of equilibrium while approaching $p_{\text{CO}} = 5$ bars. Notably, the DRC for the $\text{O}_{\text{cus}} + \text{CO}_{\text{cus}}$ reaction shows a similar increase, when the CO partial pressure exceeds the oxygen partial pressure $p_{\text{O}_2} = 1$ bar by a factor of two to three. This means that the contribution X_1^α by the creation and annihilation of the desired configuration must be positive for the $\text{O}_{\text{cus}} + \text{CO}_{\text{cus}}$ reaction. This is in contrast to the case at low partial pressures, where it is negative (but close to zero). At higher pressures ($p_{\text{CO}} \in [2, 5]$ bars), the presence of CO on the surface is still critical which is reflected by the positive DRS for the CO_{cus} -adsorption. However, the removal of CO_{cus} by the $\text{O}_{\text{cus}} + \text{CO}_{\text{cus}}$ reaction is now compensated by the creation of two vacant cus-sites, which will likely be filled by CO.

Close to $p_{\text{CO}} = 5$ bars, oxygen still dominates the surface and getting CO onto the surface is still crucial, but we have

already some kind of mixed composition with both CO and oxygen on the surface. We now have some CO on the bridge sites, and, consequently, the fast $O_{\text{cus}} + \text{CO}_{\text{br}}$ reaction starts contributing to the overall TOF (compare the [supplementary material](#)). Oxygen is very strongly bound to the bridge sites and therefore covers these. The CO adsorption on the bridge site has a positive DRS for two reasons: (i) it enables the $O_{\text{cus}} + \text{CO}_{\text{br}}$ reaction and (ii) it prevents oxygen to adsorb on the pair of vacant sites created by the latter reaction. The corresponding CO_{br} -desorption plays no role, because the $O_{\text{cus}} + \text{CO}_{\text{br}}$ reaction is significantly faster and removes most of the CO_{br} . The $O_{\text{cus}} + \text{O}_{\text{br}}$ adsorption removes vacant sites, which are needed for the CO-adsorption on both types of sites, resulting in a negative DRS. The $\text{O}_{\text{br}} + \text{CO}_{\text{cus}}$ reaction has two effects. First it slightly contributes to the overall TOF and second it creates free bridge and cus sites, which both will likely be filled with CO. The $O_{\text{cus}} + \text{CO}_{\text{br}}$ reaction itself has no impact because it is so fast that no other mechanism, which would destroy the $O_{\text{cus}}/\text{CO}_{\text{br}}$ pair, can compete with it.

Increasing the CO partial pressure beyond 5 bars, the DRS and DRC of all species drop because we are moving away from the point with the highest correlation and large fluctuations. Further, we are increasing the CO content on the surface, so that it becomes less and less important to get CO onto the surface. At around 6–7 bars, we reach a point, where all DRS and DRC are close to zero except for the $O_{\text{cus}} + \text{CO}_{\text{cus}}$ reaction. As this is still the dominant contribution to the overall TOF, we can interpret this point as the best mixed phase, where the appearance of appreciable configurations is largely unaffected by accelerating one of the elementary steps. A further increase leads then to a CO dominated surface, where the DRC identifies the O_{cus} -adsorption as the rate-determining step. Having a closer look at the results for the DRS, this result is easily understood. On this CO covered surface, the crucial point is to get oxygen onto the surface. Therefore the CO_{cus} -desorption has a positive impact, as it creates empty sites. Oxygen adsorption onto cus-sites competes with CO adsorption onto the same sites, i.e., the oxygen adsorption has a positive DRS and the CO adsorption has a negative DRS. At high CO partial pressures, CO ad/desorption is equilibrated and their DRS must cancel when summing to the DRC. Once on the surface, oxygen binds very strongly and is almost exclusively removed by forming CO_2 , with the dominant reaction being $O_{\text{cus}} + \text{CO}_{\text{br}}$ and smaller contributions by the $O_{\text{cus}} + \text{CO}_{\text{cus}}$ reaction. Due to the stronger binding of CO to bridge sites (compared to cus-sites), CO_{br} -desorption is comparatively slow, so that vacancy pairs with one bridge site are rare and thus the competition between oxygen adsorption involving a bridge site and the CO_{br} -adsorption does not play a big role.

The above discussion highlights why it is beneficial to consider both DRC and DRS. At high CO partial pressures, the DRC identifies the oxygen adsorption onto cus-sites as the only rate-determining step. As the barrier height for this reaction is already at its minimum, we would get no hint how to increase the reactivity. The DRS provides this missing information. Lowering the CO partial pressure (lowering CO adsorption) or increasing the O_2 partial pressure (increasing oxygen adsorption) are two options for a better performance. Alternatively, we could modify the catalyst such that CO has a weaker

binding to the surface (increasing CO desorption). But most prominently, we find that changing the oxygen binding has no effect, as oxygen desorption is unimportant. Thus just trying to reduce CO binding is a promising route. In contrast, binding of both is important at low CO partial pressures. If we want to modify the catalyst, such that the barrier of the rate-determining $O_{\text{cus}} + \text{CO}_{\text{cus}}$ reaction is decreased, we must therefore take care not to increase oxygen binding or to reduce CO binding. On the other hand, we would not be able to tell whether CO ad/desorption is rate-determining at intermediate CO partial pressures, if we would solely consider the DRS curves in Fig. 6. Even when considering the frequencies (or the bounds in Fig. 8) for forward and backward reactions, this step looks always equilibrated. The DRC curve in Fig. 7, however, tells that this step is rate-determining and is therefore not equilibrated.

Finally, we want to draw our attention to the question, which information could be extracted from only considering the bounds in Fig. 8 as sensitivity measures. As expected, the bounds overestimate the DRS, often by more than one order of magnitude. On the plus side, the five most important steps are correctly identified for low CO pressures. Also the steep increase of the most absolute DRS is properly reflected by the bounds. The O_{cus} desorption does not follow this trend and this is properly reproduced by the corresponding DRS bound in Fig. 8. For high CO pressures, the adsorption and desorption of CO on cus-sites are ranked most important, but the oxygen adsorption on cus-sites is only ranked seventh according to the Fisher Information Metric (FIM) based criterion. Furthermore, a number of processes involving bridge sites are identified as important. We would therefore miss the important feature that only the kinetics on the cus sites determines the reaction rate for most of the considered reaction conditions. So, a FIM based sensitivity analysis²⁷ provides a first, more global insight into the reaction kinetics, but a detailed picture of the microscopic reaction pathways requires a deeper analysis.

VI. CONCLUSION

We have presented a three-stage procedure to determine local sensitivity indices, specifically the degree of rate sensitivity, from which other sensitivity measures as the degree of rate control can easily be derived. In the first stage, bounds for the DRSs are obtained to eliminate those DRSs which are close to zero. In the second stage, the remaining DRSs are estimated using a direct sampling approach, which is based on a truncated series expansion of time-integrated linear response functions. Only those DRSs, which cannot be accurately enough estimated in the previous stages, are finally obtained using coupled finite differences. The devised approach can lead to significant computational savings, compared with the straightforward approach solely employing numerical derivatives.

We have demonstrated the approach on the example of the CO oxidation on $\text{RuO}_2(110)$. Comparing with our previous study for the sensitivity analysis of the same model,¹⁶ we arrive at significantly more accurate estimates for the sensitivity indices in a fraction of the CPU-time. For this model,

we have also demonstrated that the information obtained during the three-stage procedure can be used to derive a detailed interpretation of the nature of the microscopic reaction kinetics. We found that the DRS criterion is a suitable complement to the more common DRC and allows to explain why a step is rate-determining.

The methods for sampling the bounds and the IRF approach have been implemented into the open-source kMC package kmos and we are currently preparing it to become a part of the next publicly available update. The CFD has been implemented using just the front-end and future development will target at allowing for a more efficient implementation but also at alternative coupling strategies.⁷ Also some effort needs to be directed towards a more balanced strategy minimizing the overall computational costs.

SUPPLEMENTARY MATERIAL

See [supplementary material](#) for additional information on the derivations of series expansions of the TAC and the IRF and the pseudocode for their sampling as well as for the DRS sum rule.

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