# Tail Estimates for the Space Complexity of Randomized Incremental Algorithms ${ }^{\circ}$ 

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#### Abstract

We give tail estimates for the space complexity of randomized incremental algorithms for line segment intersection in the plane. For $n$ the number of segments, $m$ is the number of intersections, and $m \geq n \ln n \ln ^{(3)} n$, there is a constant $c$ such that the probability that the total space cost exceeds $c$ times the expected space cost is $e^{-\Omega(m /(n \ln n))}$.


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## 1 Introduction

Randomized incremental algorithms have received considerable attention recently; cf. [3], [10] and [1]. They solve a large number of geometric problems, including the construction of Voronoi diagrams and convex hulls and intersection of line segments, in optimal expected time and space. Mehlhorn [9] has given a tail estimate for the space complexity of some of these algorithms. In particular, he has shown that, for the construction of planar Voronoi diagrams and Delaunay triangulations (using the algorithm of [5]), the probability that the space cost exceeds $c$ times the expected value is at most $(c / e)^{-c} / e$ for every constant $c \geq 1$. In this paper we obtain a considerably sharper bound for line segment intersection which is exponential in the size of the problem.

This paper is structured as follows. In Section 2 we prove a simple probabilistic lemma which we then apply in Section 3 to derive the tail estimate mentioned above.

## 2 A Probalistic Lemma

For functions $M: \mathbb{N} \rightarrow \mathbb{N}$ and $d: \mathbb{N} \rightarrow \mathbb{N}$ and integers $n$ and $r$ with $n \geq r \geq 0$, call a rooted tree $T$ an $(n, r)$-tree respecting $M$ and $d$ iff either $r=0$ and $T$ consists of a single node, or $r>0$, the root of $T$ has $n$ subtrees each of which is an $(n-1, r-1)$ tree respecting $M$ and $d$, and the $n$ edges incident to the root are labeled with non-negative weights $d_{i}$ so that $d_{i} \leq d(n)$ for $1 \leq i \leq n$ and $\sum_{1 \leq i \leq n} d_{i} \leq M(n)$.

For a path $\pi$ in $T$, let $X=X_{\pi}$ be the sum of the weights of the edges along the path. The uniform distribution on the $n(n-1) \cdots(n-r+1)$ paths in $T$ makes $X$ a random variable with expectation $\mathbf{E}(X) \leq \sum_{0 \leq i \leq r-1} M(n-i) /(n-i)$.

Lemma 1 For all $t>0$ and $B \geq 0$

$$
\operatorname{Prob}(X \geq B) \leq e^{-t B+\sum_{0 \leq i \leq r-1} \frac{M(n-i)}{(n-i) d(n-i)}\left(e^{t d(n-i)}-1\right)}
$$

Remark: The proof is an adaption of the standard proof for Hoeffding's inequality, cf. [8]. The case $r=n$ and $d(i)=M(i)$ for all $i$ was previously treated in [9].

Proof. For $0 \leq i \leq r-1$, let $X_{i}$ be the weight of the $(i+1)$-st edge on the path $\pi$. Then $X=\sum_{0 \leq i \leq r-1} X_{i}$ and

$$
\operatorname{Prob}(X \geq B)=e^{-t B} e^{t B} \operatorname{Prob}\left(e^{t X} \geq e^{t B}\right) \leq e^{-t B} \mathbf{E}\left(e^{t X}\right) \leq e^{-t B} \mathbf{E}\left(\prod_{0 \leq i \leq r-1} e^{t X_{i}}\right)
$$

Put $Z_{i}=e^{t X_{i}}$ and $Q_{i}=Z_{0} Z_{1} \cdots Z_{i}$. Then, for $i \geq 1$,

$$
\mathbf{E}\left(Q_{i}\right)=\mathbf{E}\left(Q_{i-1} Z_{i}\right)=\mathbf{E}\left(Q_{i-1} \mathbf{E}\left(Z_{i} \mid Q_{i-1}\right)\right)
$$

But

$$
\mathbf{E}\left(Z_{i} \mid Q_{i-1}\right)=\frac{1}{n-i}\left(e^{t d_{1}}+\cdots+e^{t d_{n-i}}\right)
$$

where $d_{1}, d_{2}, \cdots, d_{n-i}$ are the weights of the edges emanating from the node corresponding to $Q_{i-1}$. Since $d_{j} \leq d(n-i)$ for $1 \leq j \leq n-i$ and $\sum_{1 \leq j \leq n-i} d_{j} \leq M(n-i)$, the last expression is maximized when $\lfloor M(n-i) / d(n-i)\rfloor$ weights $d_{j}$ are equal to $d(n-i)$, one weight is equal to $M(n-i)-\lfloor M(n-i) / d(n-i)\rfloor d(n-i)$ and the remaining weights are equal to zero. It follows that

$$
\begin{gathered}
\mathbf{E}\left(Z_{i} \mid Q_{i-1}\right) \leq \frac{\frac{M(n-i)}{d(n-i)} e^{t d(n-i)}+n-i-\frac{M(n-i)}{d(n-i)}}{n-i} \\
\quad=1+\frac{M(n-i)}{(n-i) d(n-i)}\left(e^{t d(n-i)}-1\right) .
\end{gathered}
$$

the case $i=0$ is treated similarly, except that no conditional probability is used. Hence,

$$
\mathbf{E}\left(\prod_{0 \leq i \leq r-1} e^{t X_{i}}\right) \leq e^{\sum_{0 \leq i \leq r-1} \frac{M(n-i)}{(n-i)(n-i)}\left(e^{t d(n-i)}-1\right)}
$$

since $\prod_{0 \leq i \leq r-1}\left(1+y_{i}\right) \leq \prod_{0 \leq i \leq r-1} e^{y_{i}}=e^{\sum_{0 \leq i \leq r-1} y_{i}}$ for all reals $y_{0}, \cdots, y_{r-1}$.
Remark: If $d(\cdot)$ is a nondecreasing function, then

$$
\frac{e^{t d(n-i)}-1}{d(n-i)} \leq \frac{e^{t d(n)}-1}{d(n)}
$$

This simplifies Lemma 1 to

$$
\operatorname{Prob}(X \geq B) \leq e^{-t B+\left(\left(e^{t d(n)}-1\right) / d(n)\right) \cdot \sum_{0 \leq i \leq r-1} \frac{M(n-i)}{n-i}}
$$

In what follows we will use this revised bound.

## 3 The Space Complexity of Randomized Incremental Constructions

Randomized incremental constructions take a random permutation of some set $S$ of $n$ objects, e.g., points in the plane, and construct a set of so-called regions, e.g., regions of the Voronoi diagram, in an incremental fashion. The algorithm maintains the collection of regions for the current subset of objects, and updates it after each insertion of a new object. The space complexity of the algorithm is defined to be the overall number of regions created during the random insertion process. This notation is appropriate because one of the main techniques for randomized incremental constructions maintains all regions ever constructed "on top of each other" without erasing old regions; see $[1,2,5,12,11]$.

Whenever an object is added, an incremental space cost occurs. In most applications this space cost depends only on the object $x$ added and the set $R$ of objects previously added (i.e. it is independent of the order in which the elements in $R$ have been inserted). For $R \subseteq S$ and $x \in S-R$ let $c(R, x)$ denote the incremental space cost when object $x$ is added to set $R$ of objects. For the analysis of the total space cost, we define a space cost tree $T_{0}$ as follows. In tree $T_{0}$ the nodes of depth $i, 0 \leq i \leq n$, have exactly $n-i$ children. The nodes of depth $i$ correspond to the subsets of $S$ of size $n-i$ in a natural way. The root corresponds to $S$ and if a node $v$ corresponds to a subset $R \subseteq S$ then the children of $v$ correspond to the sets $R-\{x\}$, where $x$ ranges over $R$. The edge connecting the nodes corresponding to $R$ and $R-\{x\}$ is labeled by $c(R-\{x\}, x)$. For a path $\pi$ in $T_{0}$, let $X(\pi)$ be the sum of the edge labels on path $\pi$. Then $X(\pi)$ is the total space cost when the elements of $S$ are inserted in the order specified by $\pi$. Let $X=X(\pi)$ be the random variable defined by the uniform distribution on the paths in $T_{0}$. We call $X$ the space cost random variable. We note that paths in $T_{0}$ represent "backward" executions of the insertion process.

Let

$$
M(r)=\max \left\{\sum_{x \in R} c(R-\{x\}, x) ; R \subseteq S,|R|=r\right\}
$$

and

$$
d(r)=\max \{c(R-\{x\}, x) ; R \subseteq S,|R|=r, x \in R\}
$$

Then $T_{0}$ is an $(n, n)$-tree respecting $M$ and $d$ and hence Lemma 1 gives a tail estimate for the total space cost. For many randomized incremental constructions, e.g., planar Voronoi diagrams and Delaunay triangulations, trapezoidal decompositions, and convex hulls in three dimensions, one has $M(r)=d(r)=a r$ for some constant $a$. Then, as shown in [9], if we choose $t=\frac{1}{a n} \ln c$ and $r=n$ in Lemma 1, we obtain

$$
\operatorname{Prob}(X \geq c \cdot a n) \leq \frac{1}{e}\left(\frac{e}{c}\right)^{c}
$$

for all $c \geq 1$. We will now discuss an application where $d$ is much smaller than $M$.

Line Segment Intersections. Randomized incremental algorithms for line segment intersection are described in [3] and [10]. These algorithms have expected running time $O(n \log n+m)$ and expected space cost $O(n+m)$, where $n$ is the number of segments and $m$ is the number of intersections. Both algorithms maintain the trapezoidal decomposition $T(R)$ of the plane with respect to the set $R$ of segments. This decomposition is defined as the planar map obtained by drawing, from each endpoint and intersection point of the segments in $R$, a vertical segment up and down and extend it until it meets other segments of $R$. The incremental space cost $c(R, s)$ of adding a segment $s \in S-R$ to the current set $R$ of segments is the number of trapezoids in $T(R \cup\{s\})$ incident to $s$. Since each trapezoid is incident to at most four segments and since the number of trapezoids in $T(R)$ is at
most $1+3|R|+3 m(R)$, where $m(R)$ is the number of pairs of intersecting segments in $R$, we have

$$
M(R):=\sum_{s \in R} c(R-\{s\}, s) \leq 4(1+3|R|+3 m(R))
$$

Clearly, $0 \leq m(R) \leq|R|^{2} / 2$. We can therefore apply Lemma 1 with $d(r)=M(r)=$ $O\left(r+\min \left(r^{2}, m\right)\right)$ to derive a tail estimate for the total space cost, albeit with a fairly weak bound of the form:

$$
\operatorname{Prob}(X \geq c(n+m) \log n) \leq c^{-\Omega(c)}
$$

for all constants $c$. We leave it to the reader to verify this bound, which we now proceed to improve. There are two weak spots in the argument just offered. First, the expected value of $m(R)$ is only $O\left(r+m \cdot \frac{r^{2}}{n^{2}}\right.$, so the upper bound $M(r)$ is a gross overestimation for most values of $M(R)$. Second, as will be shown momentarily, $d(r)$ is always much smaller than the bound given above.

To obtain a sharper tail estimate we thus proceed as follows. We first show that $d(r)=O(r \alpha(r))$, where $\alpha$ is the functional inverse of Ackermann's function (cf. [7, page 156] for a definition). We next prove a tail estimate for (a quantity related to) $m(R)$, and then finally use this tail estimate to apply Lemma 1 with a function $M(r)$ smaller than $O\left(r+\min \left(r^{2}, m\right)\right)$.

Lemma 2 Let $R \subseteq S, s \in S-R$, and $r=|R|$. Then $c(R, s) \leq \beta r \alpha(r)$, where $\alpha$ is the functional inverse of Ackermann's function and $\beta$ is some constant independent of $r$.

Proof. Let $l$ be the line supporting $s$. Then $c(R, s) \leq 6+c(R, l)$ since every trapezoid in $T(R \cup\{s\})$ incident to $s$ but not incident to an endpoint of $s$ is also incident to $l$ in $T(R \cup\{l\})$, and each endpoint of $s$ is incident to only three trapezoids in $T(R \cup\{s\})$. In order to bound the number of trapezoids incident to $l$ and above $l$, perform a scan along $l$ and consider the list of top segments of these trapezoids. This list is an ( $r, 3$ ) Davenport-Schinzel sequence (cf. [7, page 151] for a definition) as there can be no five indices $i_{1}<i_{2}<i_{3}<i_{4}<i_{5}$ in this list such that the trapezoids at places $i_{1}, i_{3}, i_{5}$ have top segment $a$ and the trapezoids at places $i_{2}$, $i_{4}$ have top segment $b$, for $b \neq a$. Since the length of an $(r, 3)$ Davenport-Schinzel sequence is bounded by $O(r \alpha(r))$ (cf. [7, page 156]), the lemma follows.

For a line segment $s \in S$ let $\operatorname{deg}(s)$ be the number of intersections between $s$ and the other segments in $S$. For $R \subseteq S$, let $D(R)=\sum_{s \in R} \operatorname{deg}(s)$. Then, clearly, $D(R) \geq 2 m(R)$; in fact, $D(R)$ counts all intersections between segments in $R$ and segments in $S$, where intersections between two segments in $R$ are counted twice.

Lemma 3 Let $a_{1} \geq 4 e^{2}, a_{2}>0, x \geq 1$, and $1 \leq r \leq n$. Then

$$
\operatorname{Prob}\left(D(R) \geq\left(\frac{a_{1} r}{n}+\frac{a_{2}}{x}\right) m\right) \leq \min \left\{e^{-a_{1} m r / n^{2}},\left(\frac{4 e x r}{a_{2} n}\right)^{a_{2} m /(n x)}\right\}
$$

where $R$ is a random subset of $S$ of size $r$.

Proof. For $r \geq \frac{2 n}{a_{1}}$ there is nothing to prove since $D(R) \leq 2 m$ always holds. For $r<\frac{2 n}{a_{1}} \leq \frac{n}{2}$, we use Lemma 1, as follows. Consider the following $(n, r)$-tree $T$. The nodes of $T$ of depth $i, 0 \leq i \leq r$, correspond to subsets of $S$ of cardinality $i$; the correspondence is many to one. If node $v$ of $T$ corresponds to $R^{\prime} \subseteq S$ then the $n-\left|R^{\prime}\right|$ children of $v$ correspond to the sets $R^{\prime} \cup\{s\}$, where $s \in S-R^{\prime}$. Also, the edge connecting $R^{\prime}$ and $R^{\prime} \cup\{s\}$ is labeled with $\operatorname{deg}(s)$. In this way, the edge labels on a leaf-to-root path sum to $D(R)$, where $R$ is the subset of $S$ corresponding to the leaf. Also, with $d(i)=n$ and $M(i)=2 m$ the tree $T$ respects $d$ and $M$ and, by symmetry, each subset $R \subseteq S$ with $|R|=r$ corresponds to the same number of leaves of $T$. Thus

$$
\begin{aligned}
\operatorname{Prob}(D(R) \geq B) & \leq e^{-t B+\sum_{0 \leq i \leq r-1} \frac{2 m}{(n-2) n}\left(e^{t n}-1\right)} \\
& \leq e^{-t B+\frac{4 m r}{n^{2}}\left(e^{t n}-1\right)}
\end{aligned}
$$

for all $B \geq 0$ and $t \geq 0$. Put $B=m\left(\frac{a_{1} r}{n}+\frac{a_{2}}{x}\right)$ and $t=\frac{1}{n} \ln \left(\frac{B n}{4 m r}\right)$. Then

$$
\begin{aligned}
\operatorname{Prob}(D(R) \geq B) & \leq e^{\frac{B}{n}\left(1-\ln \left(\frac{B n}{4 m r}\right)\right)} \\
& \leq e^{\frac{a_{1} m r}{n^{2}}\left(1-\ln \left(\frac{a_{1}}{4}\right)\right)} \\
& \leq e^{-\frac{a_{1} m r}{n^{2}}}
\end{aligned}
$$

Similarly, we have

$$
\begin{aligned}
\operatorname{Prob}(D(R) \geq B) & \leq e^{\frac{a_{2} m}{n x}\left(1-\ln \left(\frac{a_{2} n}{1 r x}\right)\right)} \\
& =e^{\frac{a_{2} m}{n x} \ln \left(\frac{4 \in r x}{a_{2} n}\right)} .
\end{aligned}
$$

Theorem 4 There are absolute constants $C, \gamma>0$ such that $\operatorname{Prob}(X \geq C m) \leq$ $e^{-\gamma\left(\frac{m}{n \ln n}\right)}$ for $m \geq n \ln n \ln \ln \ln n$.

Proof. Put $x=\ln n$. Let $T_{0}$ be the space cost tree. Let $a_{1}=4 e^{2}, a_{2}=16 e$, and let $M(r)=8\left(1+3 r+\frac{3}{2} m\left(\frac{a_{1} r}{n}+\frac{a_{2}}{x}\right)\right)$.

For $1 \leq r \leq n$, define the random variable $Y_{r}$ on the paths of $T_{0}$ so that $Y_{r}(\pi)=1$ if $M(R)>M(r)$ for the set $R$ corresponding to the node of depth $n-r$ on path $\pi$, and 0 otherwise. Let $Y=\max _{1 \leq r \leq n} Y_{r}$.

Let $T_{1}$ be the following $(n, n)$-tree respecting $M$ and $d$, where $d(r)=\beta r \alpha(n)$ and $\beta$ is as in Lemma 2 . Let $v$ be any node of $T_{0}$, let $R$ be the set corresponding to $v$, and let $w$ be the node corresponding to $v$ in $T_{1}$. If $M(R) \leq M(|R|)$, then the labels of the edges emanating from $w$ in $T_{1}$ are identical to the labels of the edges emanating from $v$ in $T_{0}$; if $M(R)>M(|R|)$, then the labels of the edges emanating from $w$ are arbitrary, but respect $M$ and $d$. Let $X_{1}$ be the random variable defined by the sum of the edge labels along the paths in $T_{1}$.

Claim $1 \operatorname{Prob}(X \geq B) \leq \operatorname{Prob}\left(X_{1} \geq B\right)+\operatorname{Prob}(Y=1)$ for any $B \geq 0$.

Proof. For paths $\pi$ with $Y(\pi)=0$ we have $X(\pi)=X_{1}(\pi)$. Thus

$$
\begin{aligned}
\operatorname{Prob}(X \geq B) & \leq \operatorname{Prob}(Y=1)+\operatorname{Prob}(X \geq B \quad \text { and } \quad Y=0) \\
& =\operatorname{Prob}(Y=1)+\operatorname{Prob}\left(X_{1} \geq B \quad \text { and } \quad Y=0\right) \\
& \leq \operatorname{Prob}(Y=1)+\operatorname{Prob}\left(X_{1} \geq B\right)
\end{aligned}
$$

We next bound $\operatorname{Prob}(Y=1)$, proceeding in two steps. First we use a standard "power-of-two-trick" to show that we may essentially concentrate on $Y_{r}$ for $r$ being a power of two and then use Lemma 3 to bound the probability that $Y_{r}$ is 1 .

Claim 2 Let $\pi$ be a path in $T_{0}$. If $Y(\pi)=1$ then there is an $r=2^{l}$ for some $l$ such that $M(R)>M(r) / 2$ where $R$ is the set corresponding to the node of depth $n-r$ on path $\pi$.

Proof. If $Y(\pi)=1$ then there is an $1 \leq r^{\prime} \leq n$, such that $M\left(R^{\prime}\right)>M\left(r^{\prime}\right)$, where $R^{\prime}$ is the set corresponding to the node of depth $n-r^{\prime}$ on $\pi$. Let $r=2^{\left\lceil\log r^{\prime}\right\rceil}$. Then $r \leq n$, since $M\left(R^{\prime}\right)<M\left(r^{\prime}\right)$ for all $r^{\prime} \geq \frac{n}{2}$. Let $R$ be the set corresponding to the node of depth $n-r$ on path $\pi$. Then $M(R) \geq M\left(R^{\prime}\right)>M\left(r^{\prime}\right) \geq \frac{M(r)}{2}$.

Claim 3 Let $R \subseteq S,|R|=r$, and $M(R) \geq M(r) / 2$. Then $D(R)>m\left(a_{1} r / n+\right.$ $\left.a_{2} / x\right)$.

Proof. We have $M(R) \leq 4(1+3|R|+3 m(R))$ and $M(r)=8\left(1+3 r+\frac{3}{2} m\left(a_{1} r / n+\right.\right.$ $\left.a_{2} / x\right)$ ). Thus $M(R) \geq M(r) / 2$ implies $m(R) \geq \frac{1}{2} m\left(a_{1} r / n+a_{2} / x\right)$. Since $D(R) \geq$ $2 m(R)$, the claim follows.

Claim 4 Let $m \geq n x$. Then

$$
\operatorname{Prob}(Y=1) \leq \log (2 x) \cdot e^{-4 e^{2} m /(n x)} .
$$

Proof. Let

$$
f(r)=\left\{\begin{array}{ll}
e^{-a_{1} m r / n^{2}} & \text { if } r \geq n / x \\
\left(\frac{4 e x r}{a_{2} n}\right)^{a_{2} m /(n x)} & \text { if } r<n / x
\end{array} .\right.
$$

Then

$$
\operatorname{Prob}(Y=1) \leq \sum_{l=0}^{\lfloor\log n\rfloor-1} f\left(2^{l}\right)
$$

according to Lemma 3, and Claims 2 and 3. Next observe that this sum can be split into two subsums, the first is

$$
\begin{aligned}
\sum_{\substack{r \in[n / x, n / 2] \\
r=2^{l} \text { for some } i}} f(r) & \leq e^{-\left(a_{1} m / n^{2}\right) \cdot(n / x)} \cdot \sum_{i=0}^{\lfloor\log (x / 2)\rfloor} e^{-\left(a_{1} m / n^{2}\right) 2^{i}} \\
& \leq e^{-a_{1} m /(n x)} \cdot \log x
\end{aligned}
$$

and the second is

$$
\begin{aligned}
\sum_{l=0}^{\lfloor\log (n / x)\rfloor} f\left(2^{l}\right) & =\left(\frac{4 e x}{a_{2} n}\right)^{a_{2} m /(n x)} \cdot \sum_{l=0}^{\lfloor\log (n / x)\rfloor}\left(2^{a_{2} m /(n x)}\right)^{l} \\
& \leq\left(\frac{4 e x}{a_{2} n}\right)^{a_{2} m /(n x)} 2^{a_{2} m /(n x) \cdot(1+\lfloor\log (n / x)\rfloor)} \\
& \leq\left(\frac{4 e x}{a_{2} n} \cdot \frac{2 n}{x}\right)^{a_{2} m /(n x)} \leq\left(\frac{1}{2}\right)^{16 e m /(n x)} \\
& \leq e^{-4 e^{2} m /(n x)}
\end{aligned}
$$

(because $\left.\left(\frac{1}{2}\right)^{4}<e^{-e}\right)$. Hence

$$
\operatorname{Prob}(Y=1) \leq \log (2 x) \cdot e^{-4 e^{2} m /(n x)}
$$

as claimed.
Claim 5 Let $x=\ln n$ and assume that $m \geq n$. Then, for all $B \geq 0$,

$$
\operatorname{Prob}\left(X_{1} \geq B\right) \leq\left(\frac{B}{3 a_{3} e m}\right)^{-B /(\beta n \alpha(n))}
$$

where $a_{3}=12 a_{2}=192 e$.
Proof. The tree $T_{1}$ respects $M$ and $d$. Since $m \geq n$, it is easily verified that $T_{1}$ also respects $\bar{M}$ and $d$ where $\bar{M}(r)=a_{3} m(r / n+1 / x)$. Thus, Lemma 1 and the remark following it imply, for any $B>0$,

$$
\begin{aligned}
\operatorname{Prob}\left(X_{1} \geq B\right) & \leq e^{-t B+\sum_{1 \leq r \leq n} \frac{a_{3} m(r / n+1 / x)}{\beta r n \alpha(n)}\left(e^{\beta t n \alpha(n)}-1\right)} \\
& \leq e^{-t B+e^{\beta t n a(n)} \frac{a_{3} m}{\beta n a(n)} \sum_{1 \leq r \leq n}\left(\frac{1}{n}+\frac{1}{x r}\right)} \\
& \leq e^{-t B+e^{\beta t n \alpha a(n)} \frac{a_{3} m}{\beta n a(n)}\left(2+\frac{\ln n}{x}\right)} \\
& \leq e^{-t B+e^{\beta t n \alpha(n)} \frac{3 a_{3} m}{\beta n \alpha(n)}} .
\end{aligned}
$$

Put $t=1 /(\beta n \alpha(n)) \ln \left(B /\left(3 a_{3} m\right)\right)$. Then

$$
\operatorname{Prob}\left(X_{1} \geq B\right) \leq e^{-\frac{B}{\beta n a(n)} \ln \left(\frac{B}{3 e a_{3} m}\right)}
$$

We can now complete the proof of Theorem 4 by combining Claims 1, 4 and 5. Since $m \geq n x \ln \ln x$, we have $-4 e^{2} m /(n x)+\ln \log (2 x) \leq-m /(n x)$, so that, by Claim 4, $\operatorname{Prob}(Y=1) \leq e^{-m /(n x)}$. Claim 5 with $B>576 e^{2} m$ implies $\operatorname{Prob}\left(X_{1} \geq\right.$ $B) \leq e^{-\Omega(m / n \alpha(n))}$. The Theorem now follows from Claim 1 .
Remark: The constants $C, \gamma$ in Theorem 4 that can be derived from our analysis are probably much too large and much too small, respectively. It would be interesting to obtain finer calibration of these constants.

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